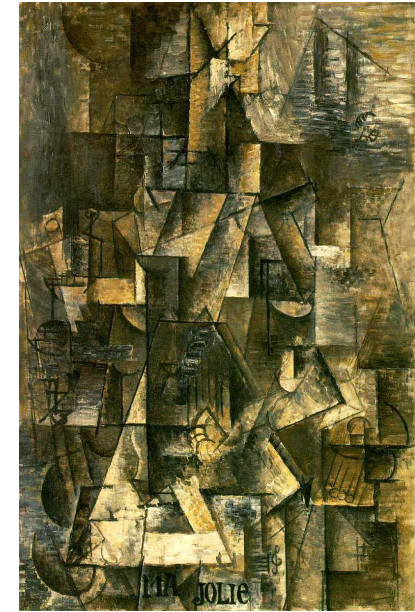
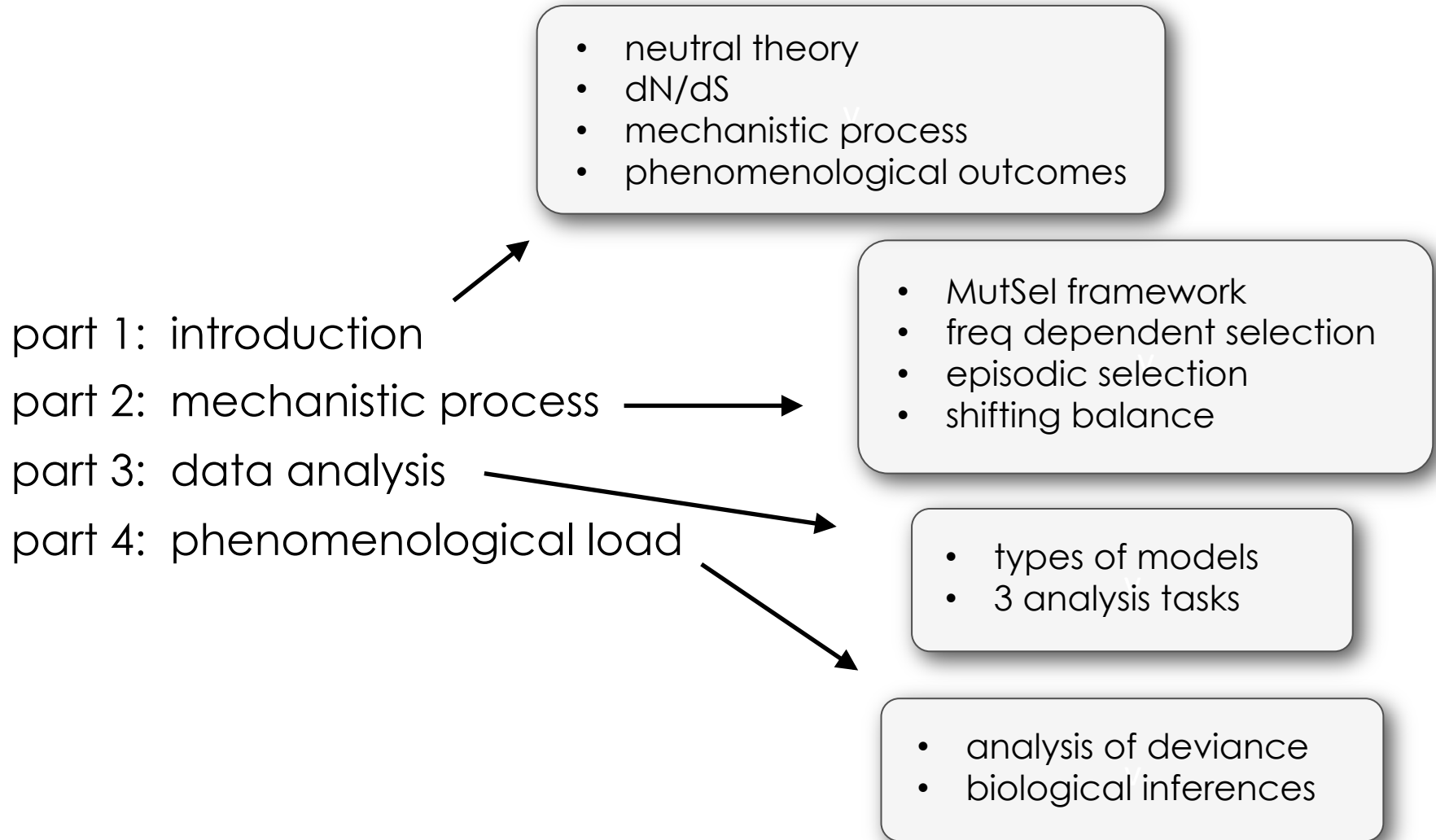


codon substitution models and the analysis of natural selection pressure

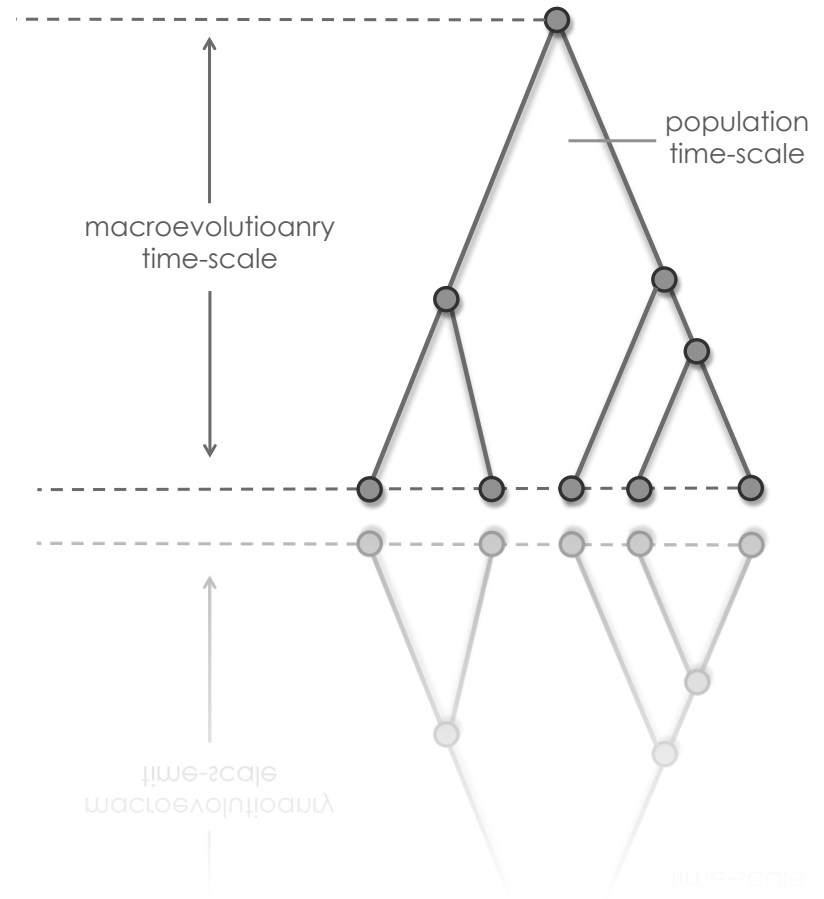
Joseph P. Bielawski
Department of Biology
Department of Mathematics & Statistics
Dalhousie University



The goals and the plan



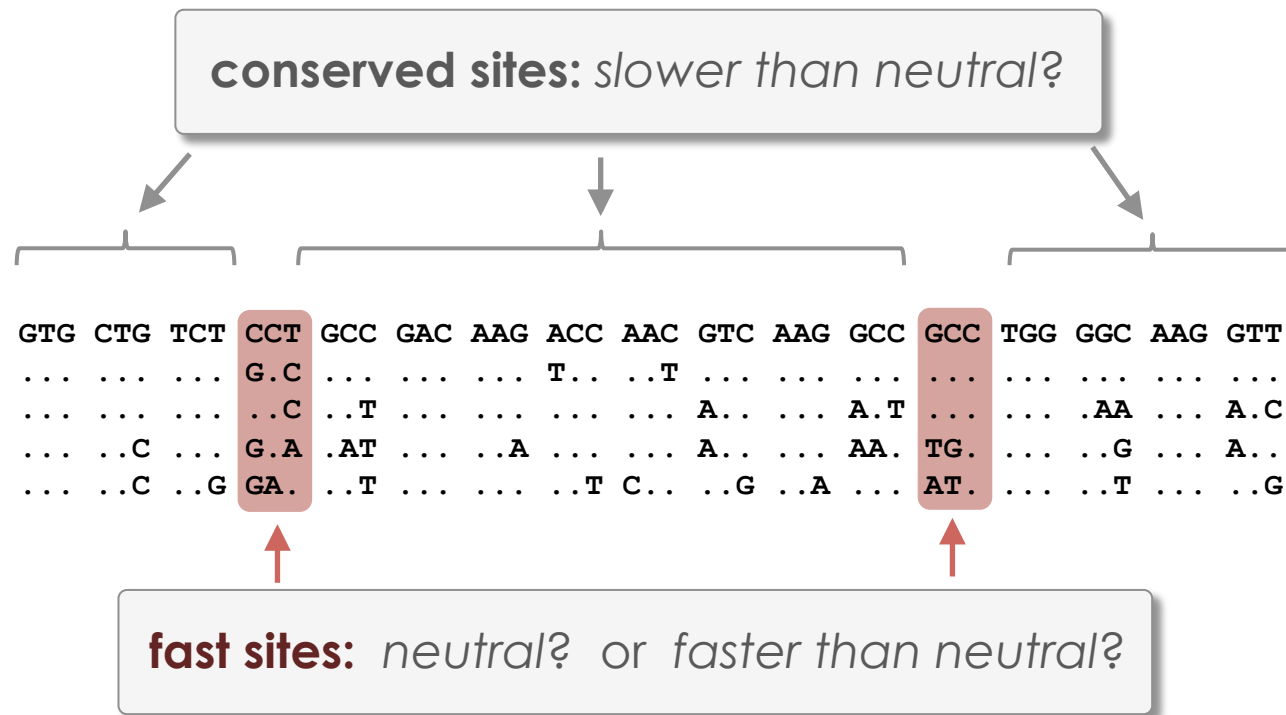
part 1: introduction



evolutionary rate depends on intensity of selection

selectively constrained = slower than neutral (drift alone)

adaptive divergence = faster than neutral (drift alone)



What is the neutral expectation?

neutral theory of molecular evolution (Kimura 1968)

the **number of new mutations** arising in a diploid population

$$2N\mu$$

the **fixation probability** of a new mutant by drift

$$1/2N$$

The **substitution (fixation) rate**, k

$$k = 2N\mu \times 1/2N$$

the elegant simplicity of **neutral theory**: $k = \mu$

genetic code determines impact of a mutation

	U	C	A	G	
U	UUU Phe UUC Phe UUA Leu UUG Leu	UCU Ser UCC Ser UCA Ser UCG Ser	UAU Tyr UAC Tyr UAA Stop UAG Stop	UGU Cys UGC Cys UGA Stop UGG Trp	U C A G
C	CUU Leu CUC Leu CUA Leu CUG Leu	CCU Pro CCC Pro CCA Pro CCG Pro	CAU His CAC His CAA Gln CAG Gln	CGU Arg CGC Arg CGA Arg CGG Arg	U C A G
A	AUU Ile AUC Ile AUA Ile AUG Met	ACU Thr ACC Thr ACA Thr ACG Thr	AAU Asn AAC Asn AAA Lys AAG Lys	AGU Ser AGC Ser AGA Arg AGG Arg	U C A G
G	GUU Val GUC Val GUA Val GUG Val	GCU Ala GCC Ala GCA Ala GCG Ala	GAU Asp GAC Asp GAA Glu GAG Glu	GGU Gly GGC Gly GGA Gly GGG Gly	U C A G

<http://www.langara.bc.ca/biology/mario/Assets/Geneticcode.jpg>

The genetic code determines how random changes to the gene brought about by the process of mutation will impact the function of the encoded protein.

Kimura (1983)

d_S : number of synonymous substitutions per synonymous site (K_S)

d_N : number of nonsynonymous substitutions per nonsynonymous site (K_A)

ω : the ratio d_N/d_S ; it measures selection at the protein level

an index of selection pressure

rate ratio	mode	example
$dN/dS < 1$	purifying (negative) selection	histones
$dN/dS = 1$	Neutral Evolution	pseudogenes
$dN/dS > 1$	Diversifying (positive) selection	MHC, Lysin

an index of selection pressure

Why use d_N and d_S ? (Why not use raw counts?)

example of counts:

300 codon gene from a pair of species

5 synonymous differences

5 nonsynonymous differences

$$5/5 = 1$$

why don't we conclude that rates are equal (i.e., **neutral evolution**)?

the genetic code & mutational opportunities

Relative proportion of different types of mutations in hypothetical protein coding sequence.				
Type	Expected number of changes (proportion)			
	All 3 Positions	1st positions	2nd positions	3rd positions
Total mutations	549 (100)	183 (100)	183 (100)	183 (100)
Synonymous	134 (25)	8 (4)	0 (0)	126 (69)
Nonsynonymous	392 (71)	166 (91)	176 (96)	57 (27)
nonsense	23 (4)	9 (5)	7 (4)	7 (4)

Modified from Li and Graur (1991). Note that we assume a hypothetical model where all codons are used equally and that all types of point mutations are equally likely.

Why do we use d_N and d_S ?

same example, but using d_N and d_S :

Synonymous sites = 25.5%

$$S = 300 \times 3 \times 25.5\% = 229.5$$

Nonsynonymous sites = 74.5%

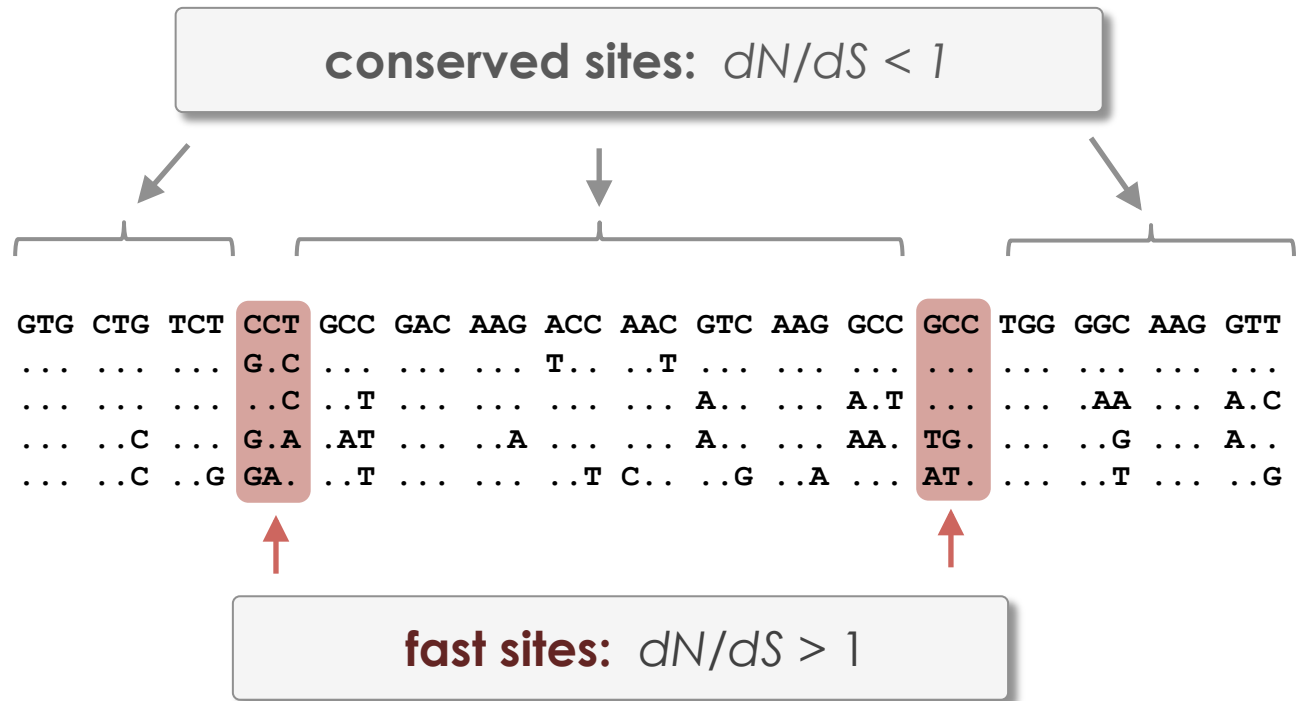
$$N = 300 \times 3 \times 74.5\% = 670.5$$

So, $d_S = 5/229.5 = 0.0218$

$$d_N = 5/670.5 = 0.0075$$

$d_N/d_S (\omega) = 0.34$, **purifying selection !!!**

an index of selection pressure acting on the protein



conclusion: dN differs from dS due to the effect of selection on the protein.

mutational opportunity vs. physical site

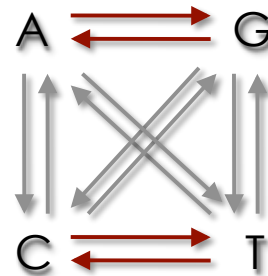
Relative proportion of different types of mutations in hypothetical protein coding sequence.				
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nonsense	23 (4)	9 (5)	7 (4)	7 (4)

Note that by framing the counting of sites in this way we are using a “mutational opportunity” definition of the sites. Thus, a synonymous or non-synonymous site is not considered a physical entity!

Note that we assume a hypothetical model where all codons are used equally and that all types of point mutations are equally likely.

real data have biases (*Drosophila GstD1* gene)

transitions vs. **transversions**:



$ts/tv = 2.71$

preferred vs. **un-preferred** codons:

partial codon usage table for the *GstD* gene of *Drosophila*

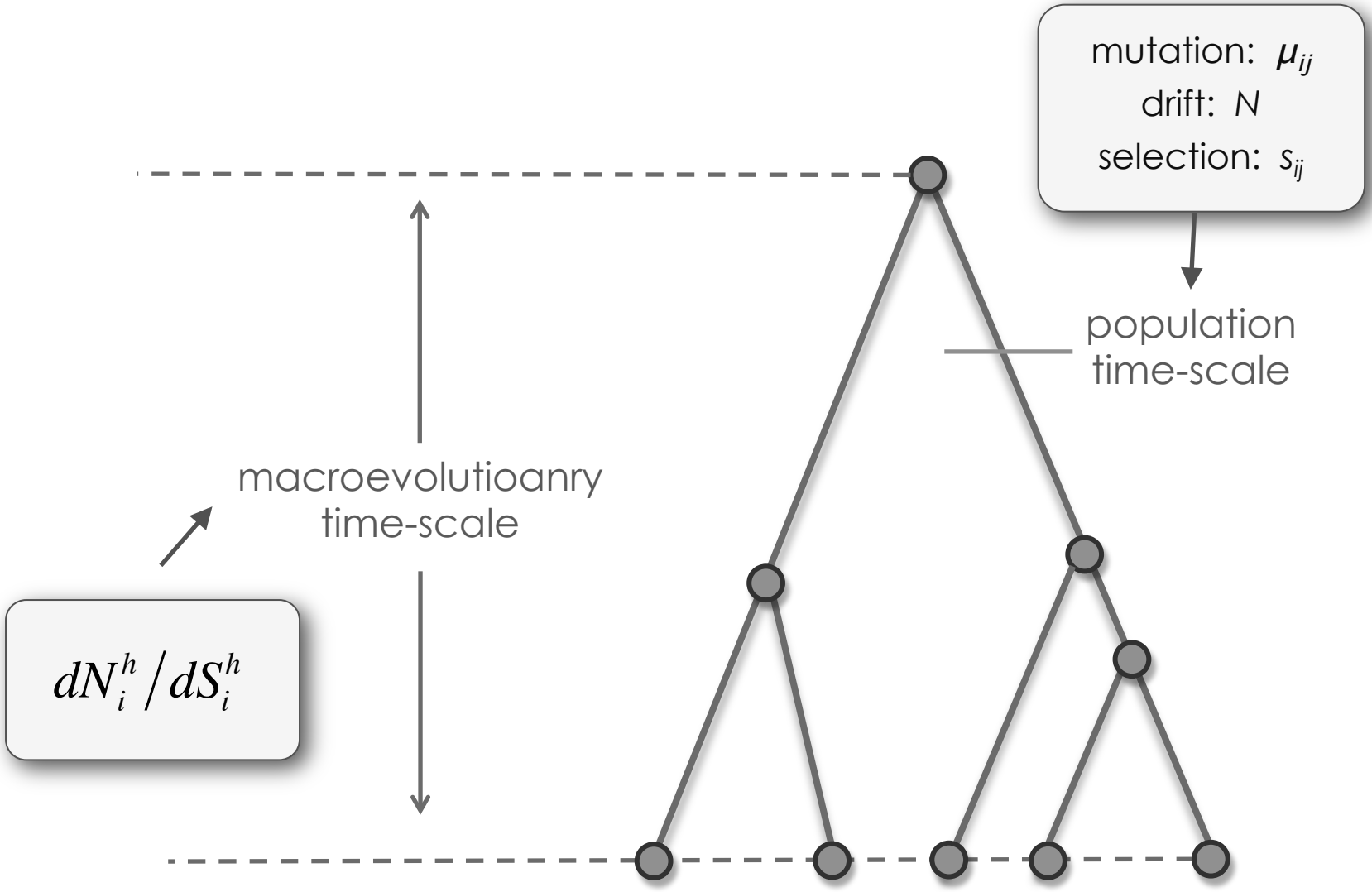
Phe F	TTT	0	Ser S	TCT	0	Tyr Y	TAT	1	Cys C	TGT	0
	TTC	27		TCC	15		TAC	22		TGC	6
Leu L	TTA	0		TCA	0	*** *	TAA	0	*** *	TGA	0
	TTG	1		TCG	1		TAG	0	Trp W	TGG	8
Leu L	CTT	2	Pro P	CCT	1	His H	CAT	0	Arg R	CGT	1
	CTC	2		CCC	15		CAC	4		CGC	7
	CTA	0		CCA	3	Gln Q	CAA	0		CGA	0
	CTG	29		CCG	1		CAG	14		CGG	0

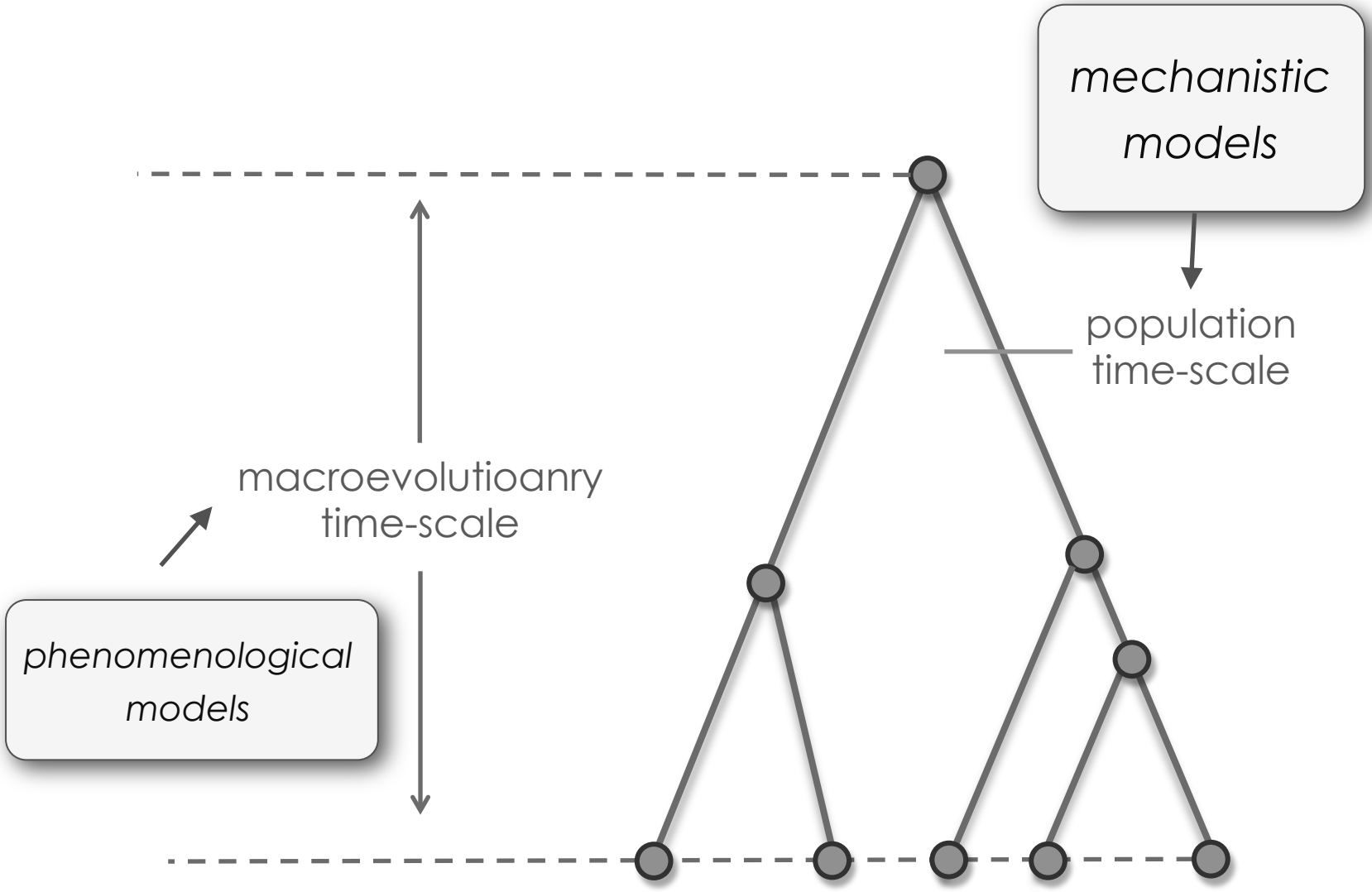
an index of selection pressure acting on the protein

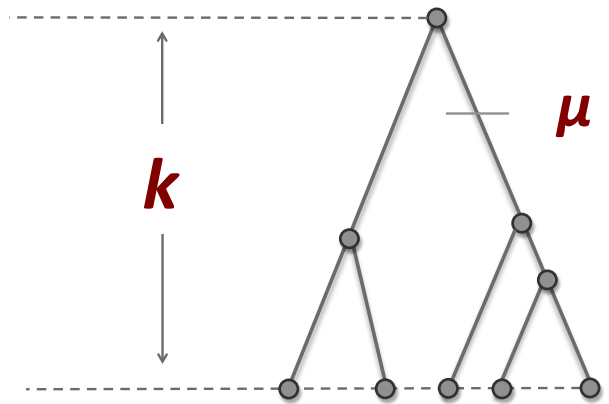
$$\omega = \frac{dN}{dS}$$

Don't worry: we will improve upon the counting method later in this lecture via likelihood!

correcting dS and dN for underlying mutational process of the DNA makes them **sensitive to assumptions about the process of evolution!**







- Wright-Fisher population
- drift: N
- mutation: μ
- selection: s_{ij}
- s_{ij} vary among sites AND amino acids
- expected dN^h/dS^h

“MUTSEL MODELS”

$$\text{Pr} = \begin{cases} \mu_{ij} & \text{if neutral} \\ \mu_{ij} N \times \frac{2s_{ij}}{1 - e^{-2Ns_{ij}}} & \text{if selected} \end{cases}$$

$$s_{ij} = \Delta f_{ij}$$

Halpern and Bruno (1998)

fixation probability with selection

population genetics at a single codon site (h)

fitness coefficients

$$f^h = \langle f_1, \dots, f_{61} \rangle$$

selection coefficients

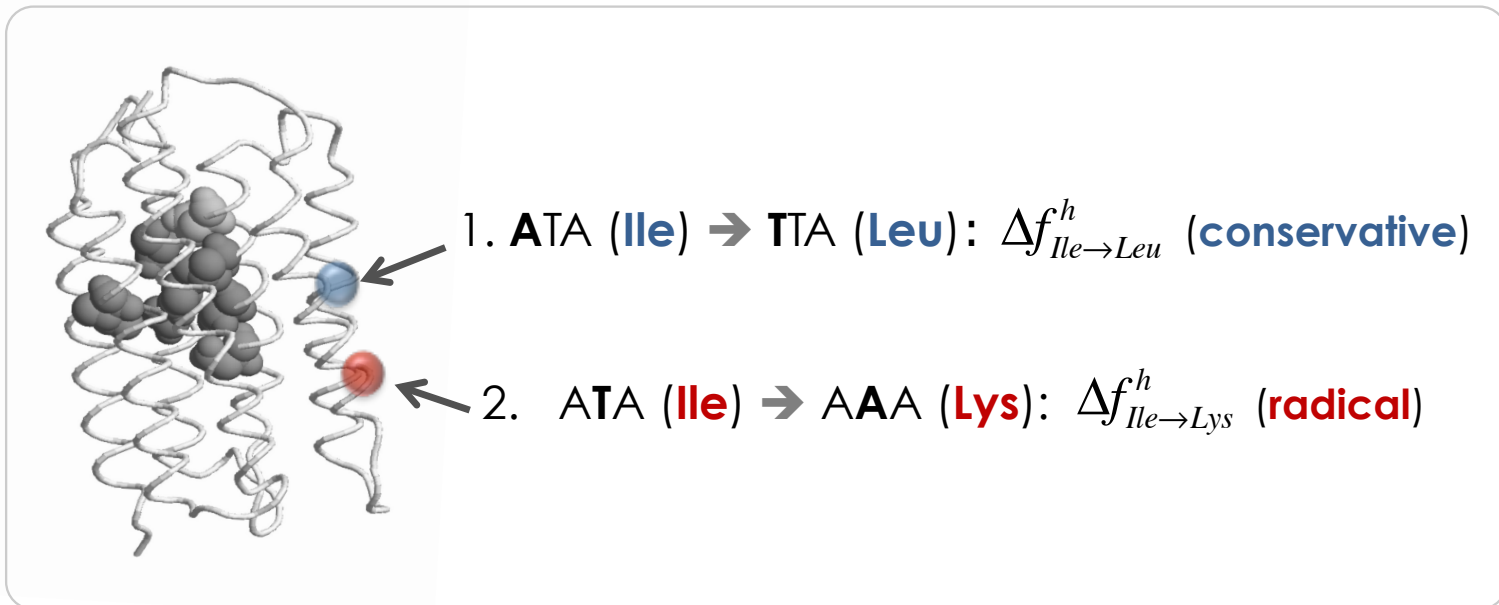
$$s_{ij}^h = f_j^h - f_i^h$$

fixation probability (Kimura, 1962)

$$\Pr(s_{ij}^h) = \frac{2s_{ij}^h}{1 - e^{-2Ns_{ij}^h}}$$

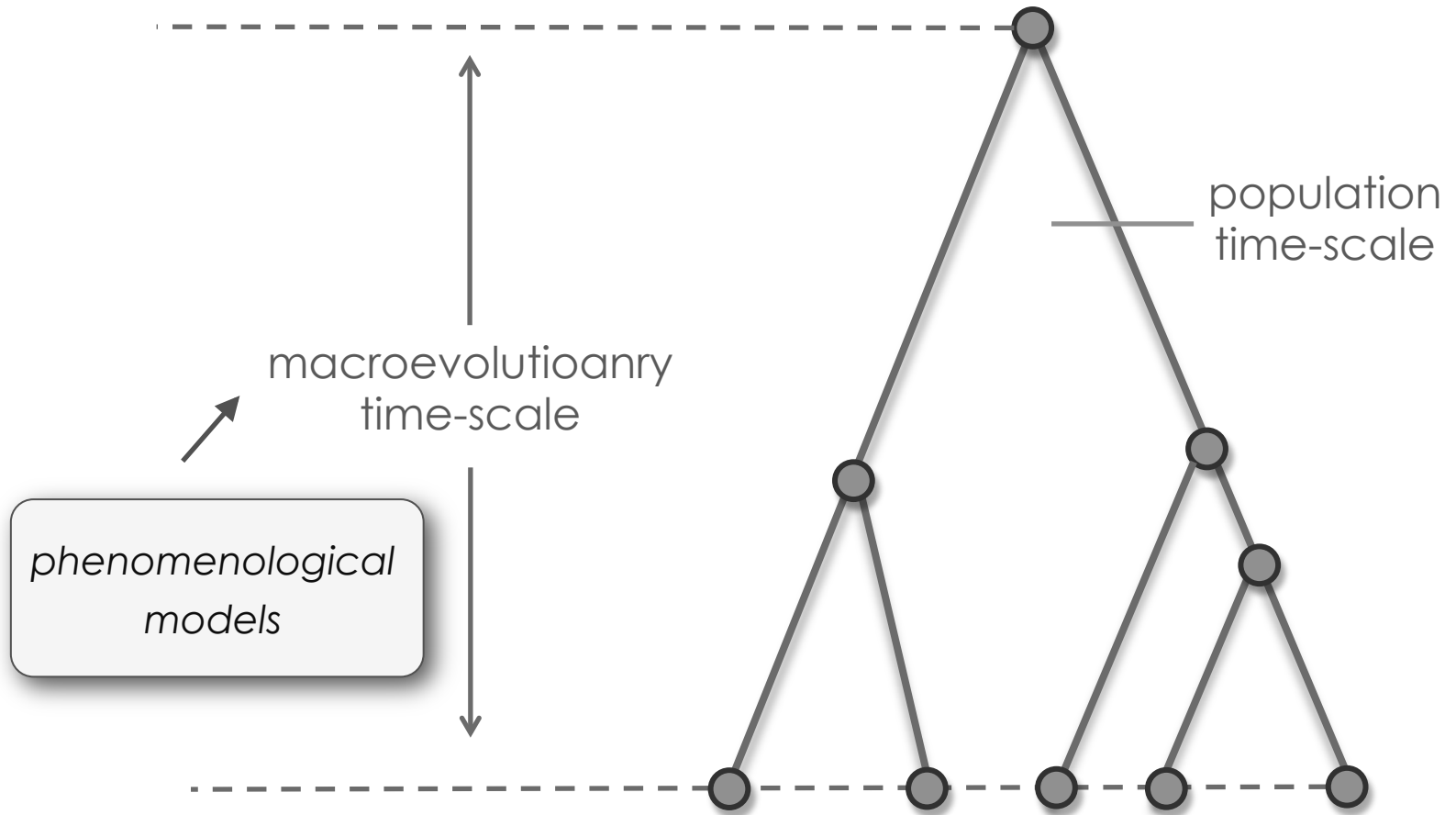
fixation probability with selection

MutSel: selection favours amino acids with higher fitness (if N is large enough)

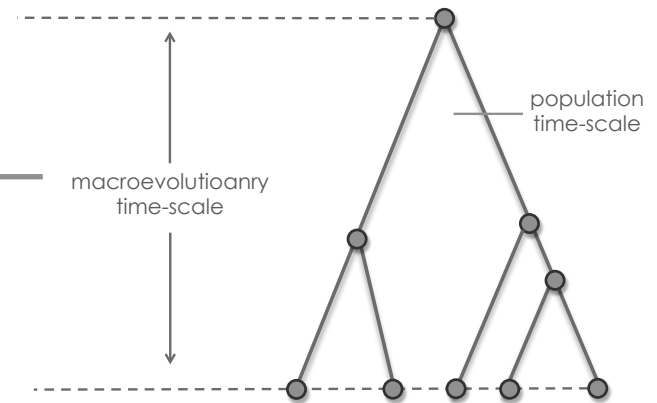


realism: fitness expected to differ among sites and amino acids according to protein function

the cost of realism: too complex to fit such a model to real data (but simplified versions will allow new ways of data analysis)



phenomenological models



“OMEGA MODELS”

$$q_{ij} = \begin{cases} 0 & \text{if } i \text{ and } j \text{ differ by } > 1 \\ \pi_j & \text{for synonymous tv.} \\ \kappa\pi_j & \text{for synonymous ts.} \\ \omega\pi_j & \text{for non-synonymous tv.} \\ \omega\kappa\pi_j & \text{for non-synonymous ts.} \end{cases}$$

Goldman and Yang (1994)
Muse and Gaut (1994)

- phenomenological parameters
- ts/tv ratio: κ
- codon frequencies: π_j
- $\omega = dN/dS$
- parameter estimation via ML
- *stationary process*

the instantaneous rate matrix, Q , is very big: 61×61

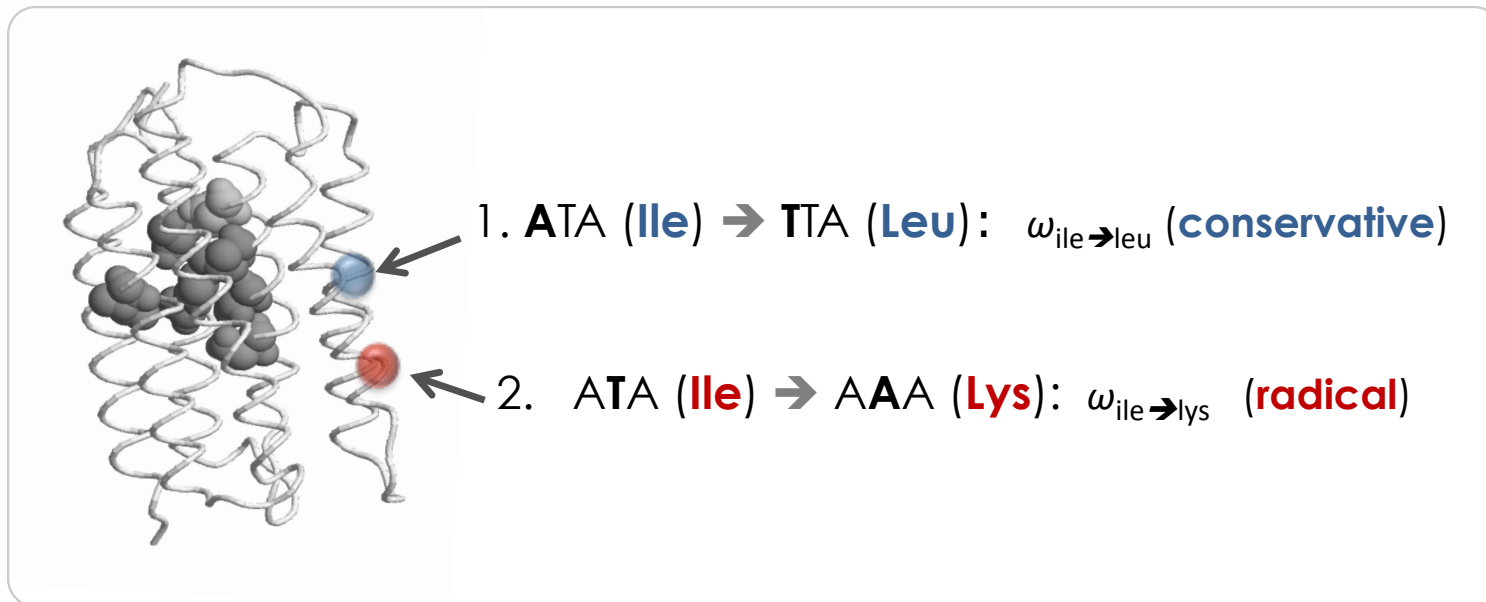
phenomenological codon models: just a few parameters are needed to cover the 3721 transitions between codons!

From codon below:	to codon below:							GGG (Gly)
	TTT (Phe)	TTC (Phe)	TTA (Leu)	TTG (Leu)	CTT (Leu)	CTC (Leu)	...	
TTT (Phe)	---	$\kappa\pi_{TTC}$	$\omega\pi_{TTA}$	$\omega\pi_{TTG}$	$\omega\kappa\pi_{TTT}$	0	...	0
TTC (Phe)	$\kappa\pi_{TTT}$	---	$\omega\pi_{TTA}$	$\omega\pi_{TTG}$	0	$\omega\kappa\pi_{CTC}$...	0
TTA (Leu)	$\omega\pi_{TTT}$	$\omega\pi_{TTC}$	---		0	0	...	0
TTG (Leu)	$\omega\pi_{TTT}$	$\omega\pi_{TTC}$	$\kappa\pi_{TTA}$	---	0	0	...	0
CTT (Leu)	$\omega\kappa\pi_{TTT}$	0	0	0	---	$\kappa\pi_{CTC}$...	0
CTC (Leu)	0	$\omega\kappa\pi_{TTC}$	0	0	$\kappa\pi_{TTT}$	---	...	0
⋮ ↓	⋮ ↓	⋮ ↓	⋮ ↓	⋮ ↓	⋮ ↓	⋮ ↓	⋮ ↓	
GGG (Gly)	0	0	0	0	0	0	0	---

* This is equivalent to the codon model of Goldman and Yang (1994). Parameter ω is the ratio d_N/d_S , κ is the transition/transversion rate ratio, and π_i is the equilibrium frequency of the target codon (i).

substitution probability with selection

intentional simplification: all amino acid substitutions have the same ω !



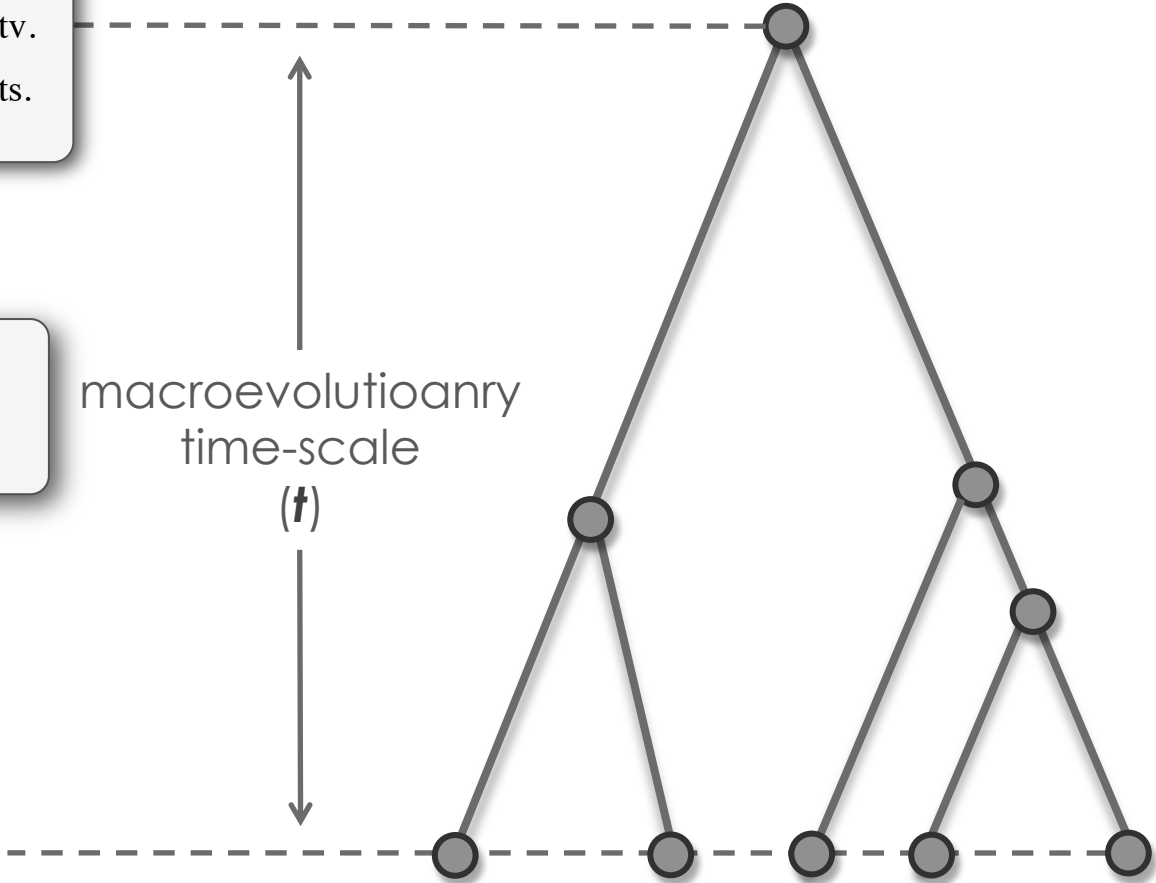
contradiction? selection should favour amino acids with higher fitness.

probability of substitution between codons over time, $P(t)$

$$Q_{ij} = \begin{cases} 0 & \text{if } i \text{ and } j \text{ differ by } > 1 \\ \pi_j & \text{for synonymous tv.} \\ \kappa\pi_j & \text{for synonymous ts.} \\ \omega\pi_j & \text{for non-synonymous tv.} \\ \omega\kappa\pi_j & \text{for non-synonymous ts.} \end{cases}$$

$$P(\mathbf{t}) = \{p_{ij}(\mathbf{t})\} = e^{Q\mathbf{t}}$$

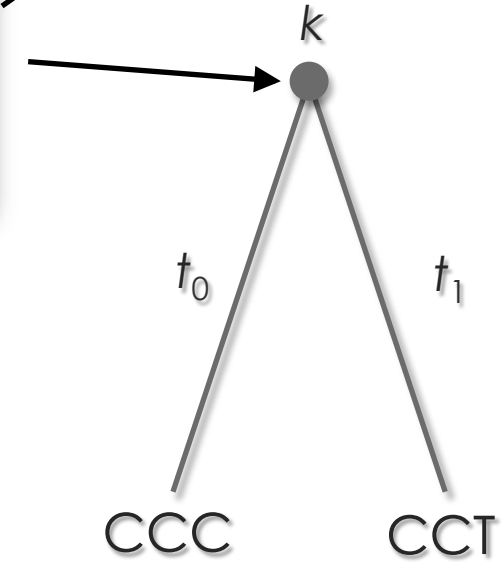
recall that **Paul Lewis** introduced **Q matrices** and how to obtain **transition probabilities**



likelihood of the data at a site

$$L_h(\text{CCC}, \text{CCT}) = \sum_k \pi_k p_{k\text{CCC}}(t_0) p_{k\text{CCT}}(t_1)$$

the likelihood is a **sum over all possible ancestral codon states** that could have been observed at node k



recall that **Paul Lewis** described how to compute the likelihood of the data at a site for a DNA model. The only difference here is that the states are codons rather than nucleotides

note: analysis is typically done by using an unrooted tree

likelihood of the data at all sites

The likelihood of observing the entire sequence alignment is the product of the probabilities at each site.

Paul Lewis covered this with the “**AND**” rule in his likelihood lecture

$$L = L_1 \times L_2 \times L_3 \times \dots \times L_N = \prod_{h=1}^N L_h$$

see **Paul Lewis's** lecture slides for more about likelihoods vs. log-likelihoods

The log likelihood is a sum over all sites.

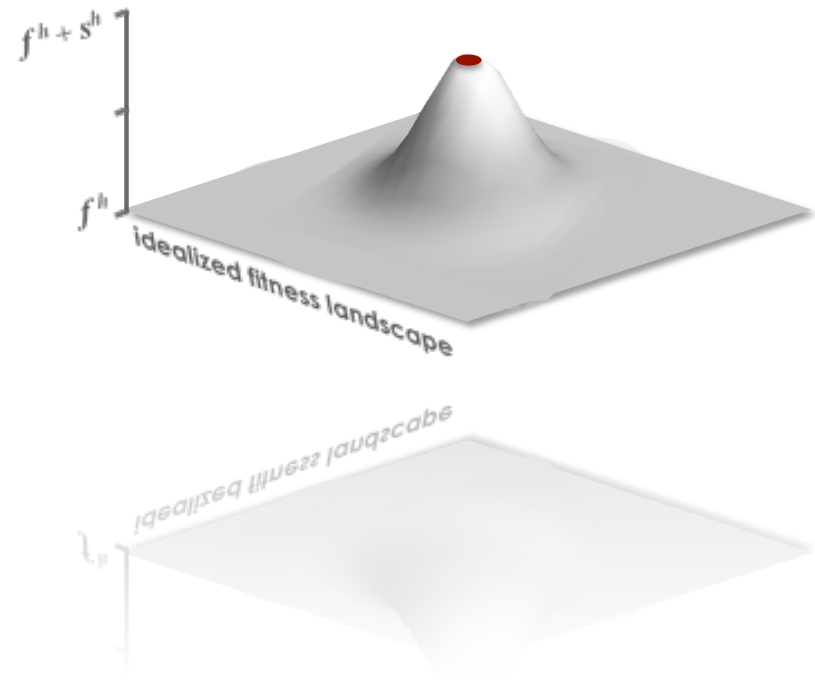
$$\ell = \ln\{L\} = \ln\{L_1\} + \ln\{L_2\} + \ln\{L_3\} + \dots + \ln\{L_N\} = \sum_{h=1}^N \ln\{L_h\}$$

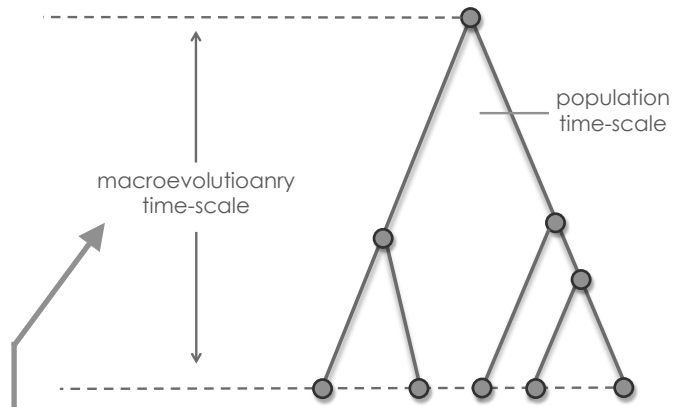
we made some progress...

1. we are now being explicit about **phenomenological and mechanistic models**
2. we are more **cautious about mechanistic interpretation** of phenomenological parameters
3. we have learned how to **connect evolutionary mechanisms to the substitution process**
4. we introduced the *idea* that we can **compute expectations** from mechanistic parameters

Lets look at some mechanism of evolution and “see” what we should expect!

part 2: mechanistic processes
of codon evolution





mechanistic models

“MUTSEL FRAMEWORK”

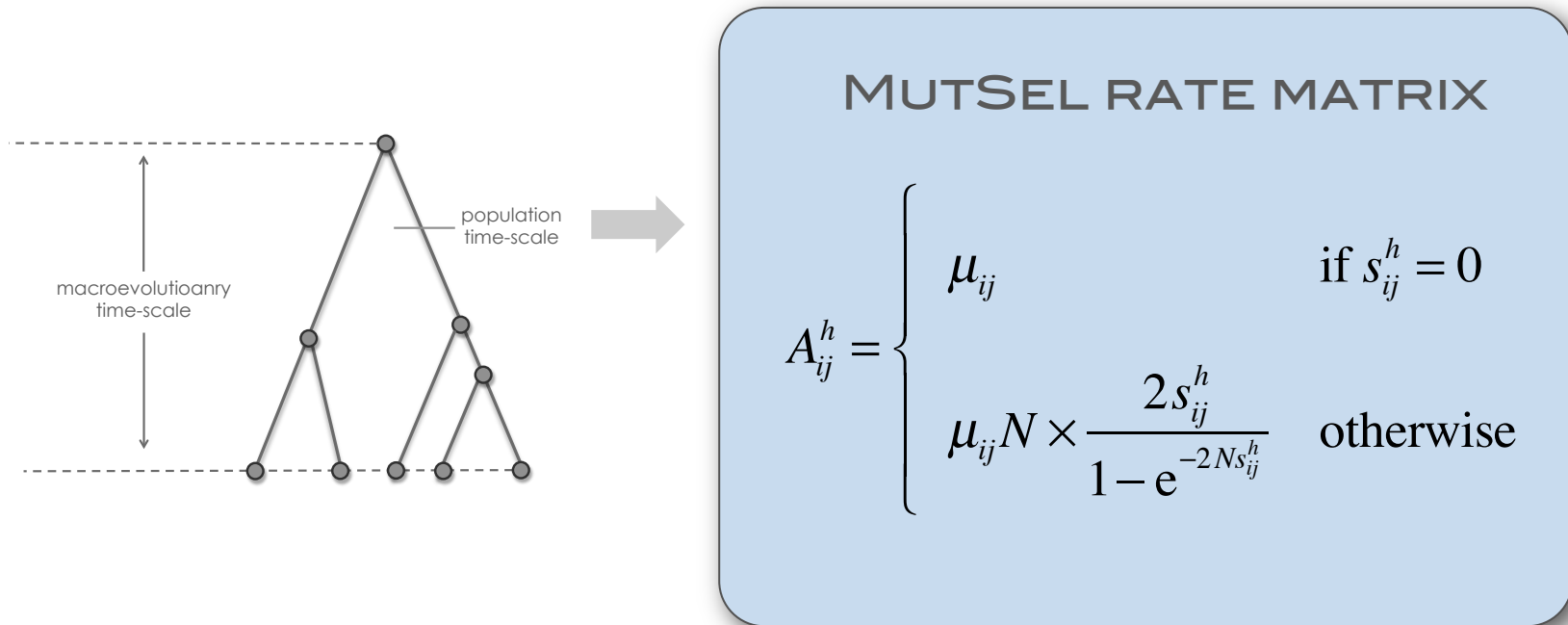
$$A_{ij}^h = \begin{cases} \mu_{ij} & \text{if } s_{ij}^h = 0 \\ \mu_{ij} N \times \frac{2s_{ij}^h}{1 - e^{-2Ns_{ij}^h}} & \text{otherwise} \end{cases}$$

$$s_{ij} = \Delta f_{ij}$$

Halpern and Bruno (1998)
Jones et al. (2016)

phenomenological models

site-specific MutSel rate matrix



- MutSel time-scale is infinitesimal compared to substitution scale
- MutSel probabilities approximate the instantaneous site-specific rate matrix, A
- μ_{ij} = nucleotide GTR process (before the effect of selection)

site-specific MutSel rate matrix

two explicit ways to reconcile **population genetics**
and **macroevolution**:

1. map fitness to equilibrium frequencies

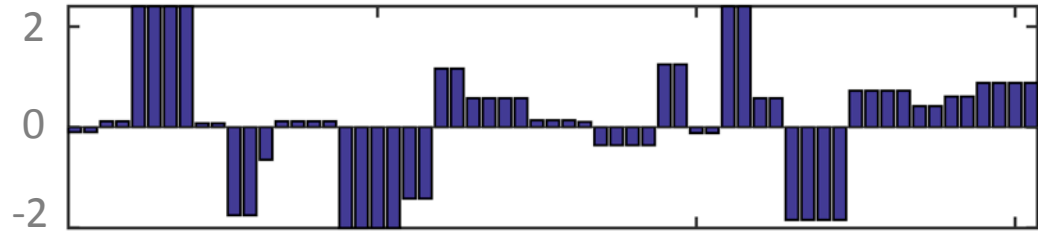
2. macroevolution index of selection intensity

(1) Sella and Hirsh 2005; (2) Jones et al. 2016

1. fitness coefficients map to stationary codon frequencies

fitness coefficients

$$f^h = \langle f_1, \dots, f_{61} \rangle$$



codon frequencies

$$\pi^h = \langle \pi_1, \dots, \pi_{61} \rangle$$



2. from fitness coefficients to dN/dS

MUTSEL RATE MATRIX

$$dN^h / dS^h = \frac{E[\text{evolution w/ selection}]}{E[\text{evolution by drift alone}]}$$

$$dN^h / dS^h = \frac{\sum_{i \neq j} \pi_i^h A_{ij}^h I_N}{\sum_{i \neq j} \pi_i^h \mu_{ij} I_N}$$

- $dN/dS = \omega$ when matrix A^h is replaced by matrix Q of model $M0$
- dN/dS is an analog of ω under MutSel

positive selection: 3 evolutionary scenarios



frequency dependent selection



episodic adaptation

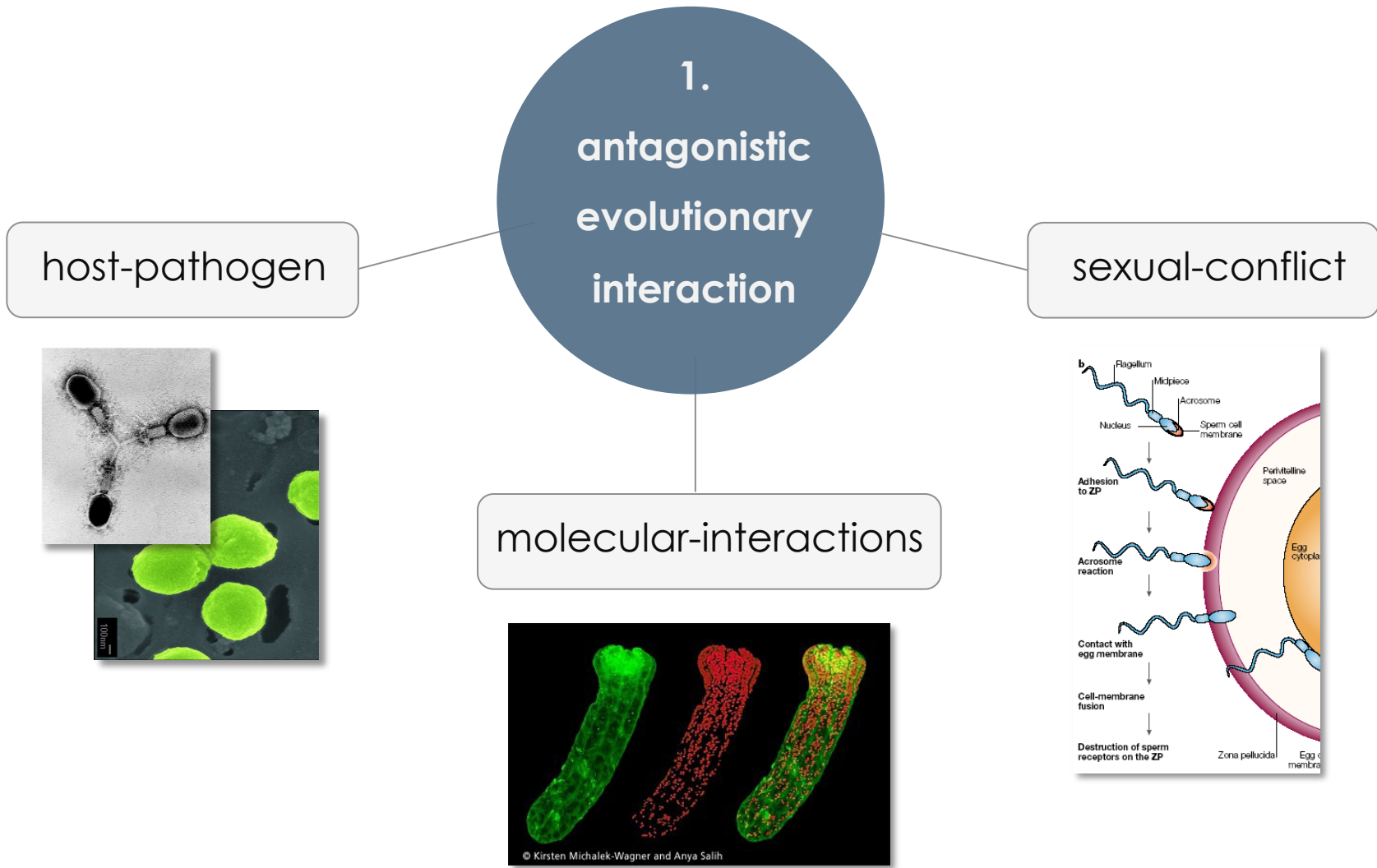


shifting balance

dynamic
fitness
landscape

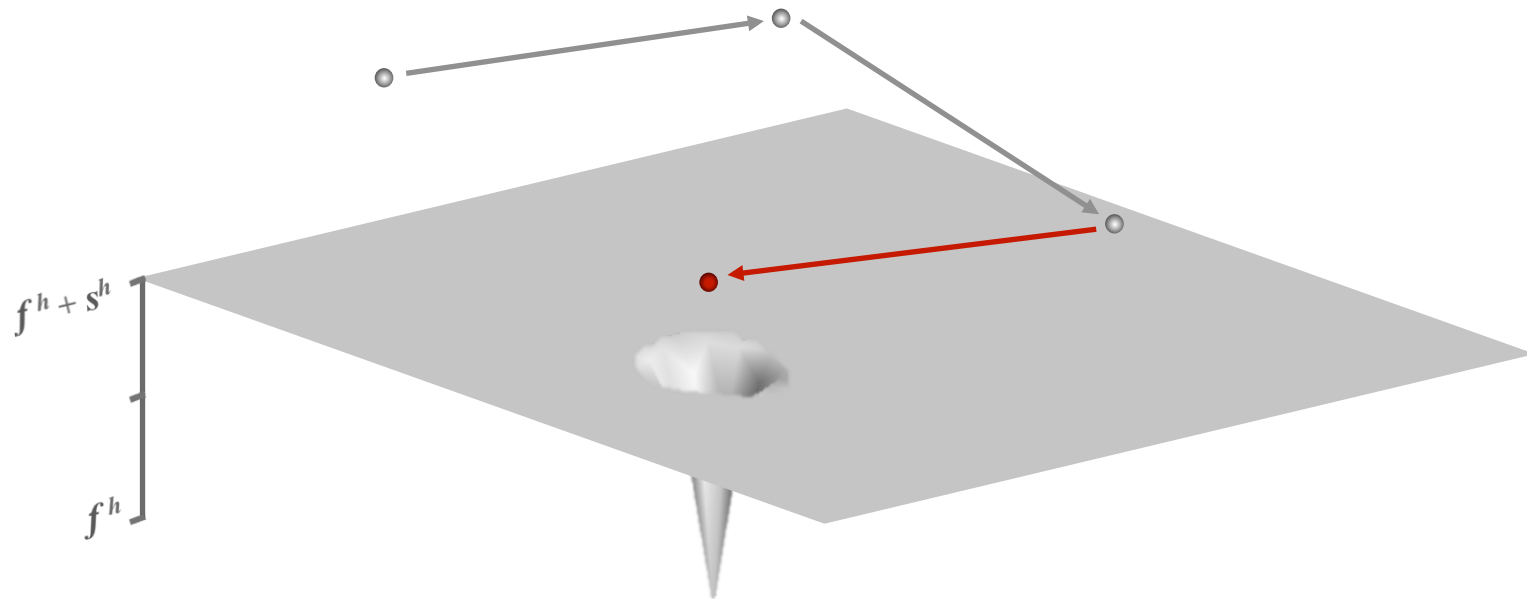
static
fitness
landscape

scenario 1: frequency dependent selection



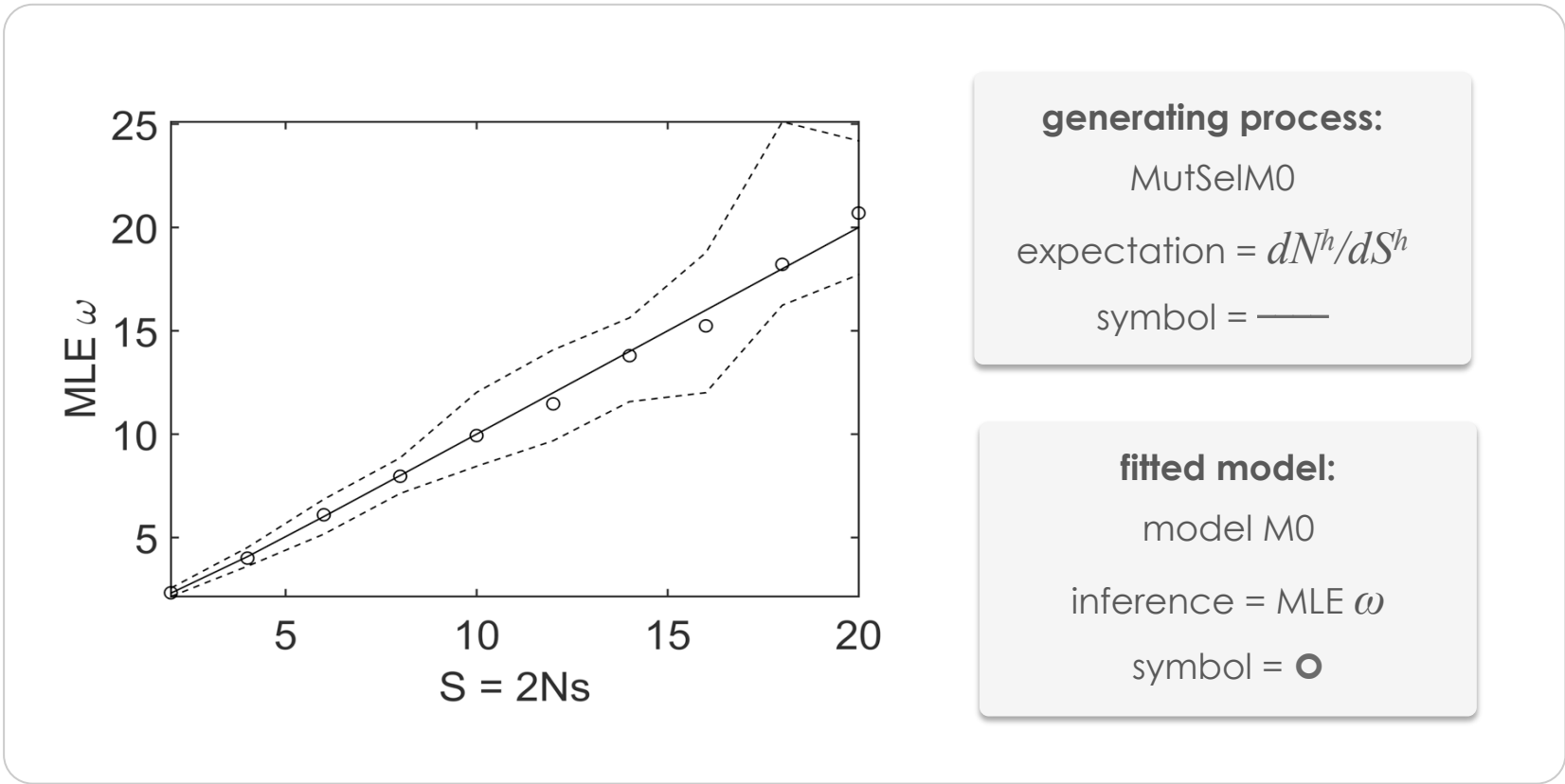
frequency-dependent selection: MutSelM0

1. amino acid at a site has f^h ; all others have $f^h + s$
2. fitness values swap when a substitution occurs



MutSelM0: (1) and (2) above imply Markov chain properties with the same rate matrix Q as **codon model M0**

frequency-dependent selection: MutSelM0



conclusion: phenomemological codon models
assume frequency-dependent selection

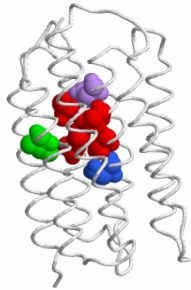
scenario 2: adaptive peak shift

2.
episodic
Darwinian
adaptation

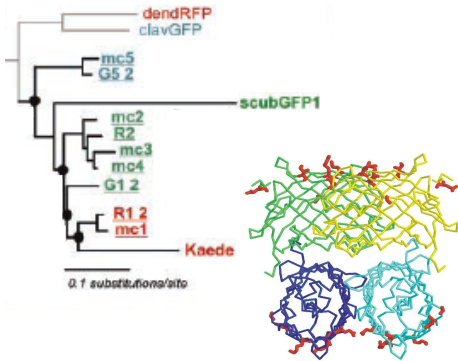
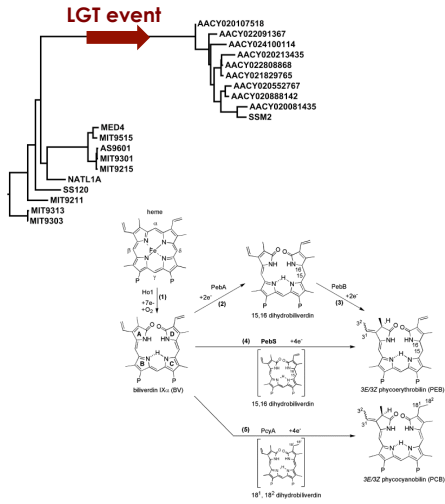
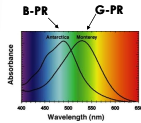
exploitation of a
new niche

lateral gene
transfer (LGT)

gene duplication

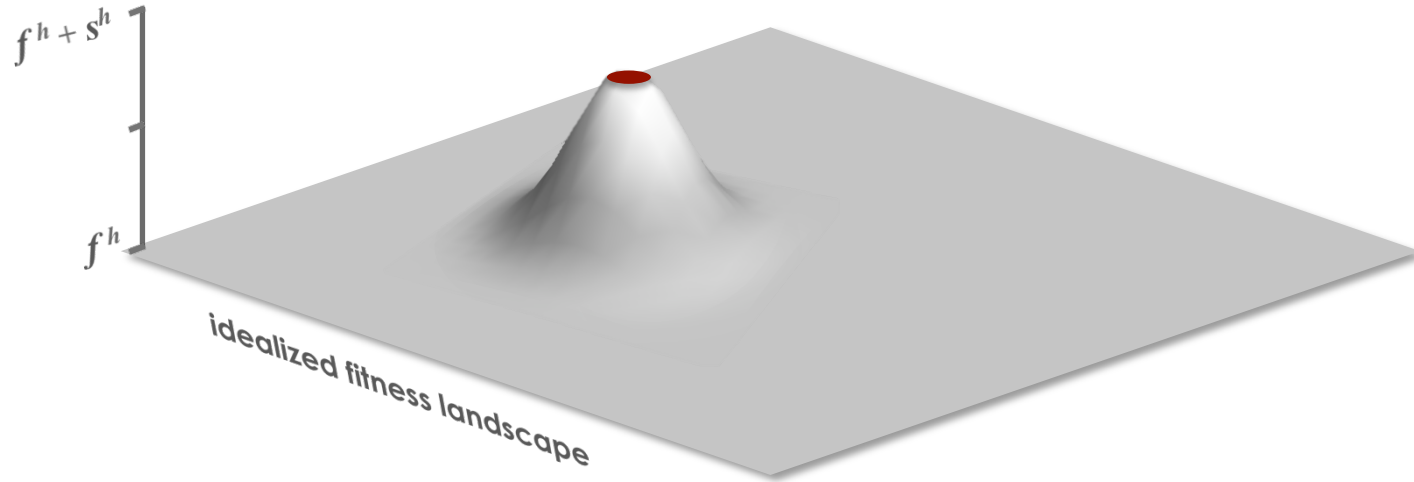


Spectral tuning switch (105)
Green (540) to Blue (490nm)



adaptive peak shift: evolution of novel function

optimal function in a stable environment



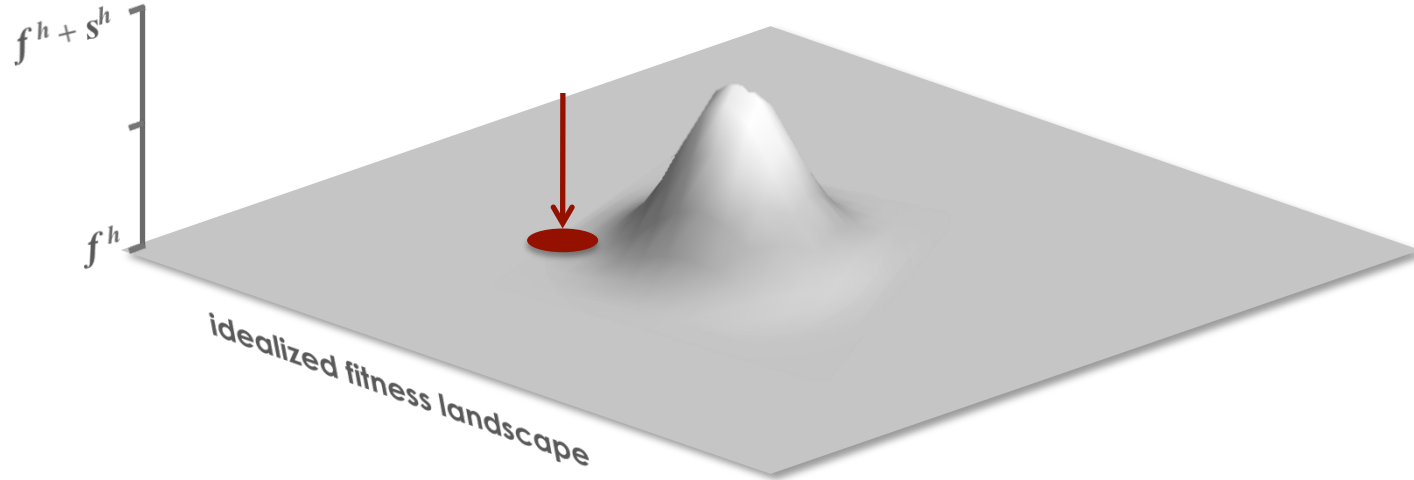
population: at fitness peak

fitness peak: stationary

FFTNS: keeps population at peak

adaptive peak shift: evolution of novel function

sub-optimal function in a novel environment



population: lower fitness

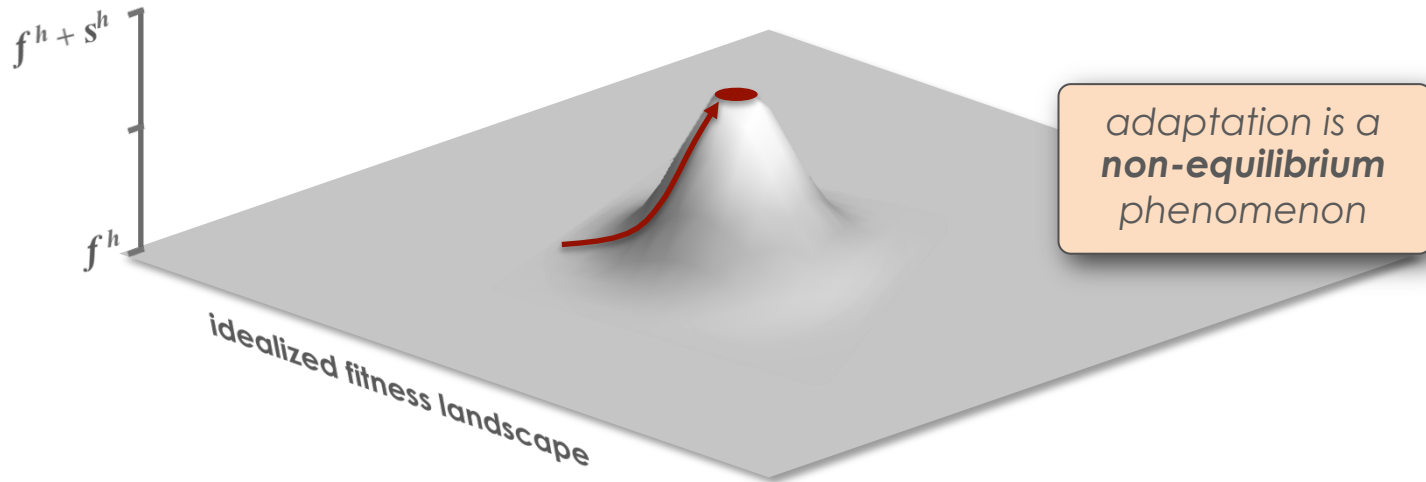
fitness peak: moving

FFTNS: increase population mean fitness

(non-stationary process)

adaptive peak shift: evolution of novel function

episodic adaptive evolution of a novel function



population: returns to peak

fitness peak: stabilized

FFTNS: increases population mean fitness until at peak

BIOLOGY LETTERS

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Research

Cite this article: dos Reis M. 2015 How to calculate the non-synonymous to synonymous rate ratio of protein-coding genes under the Fisher–Wright mutation–selection framework. *Biol. Lett.* **11**: 20141031. <http://dx.doi.org/10.1098/rsbl.2014.1031>

Received: 8 December 2014

Accepted: 16 March 2015

Molecular evolution

How to calculate the non-synonymous to synonymous rate ratio of protein-coding genes under the Fisher–Wright mutation–selection framework

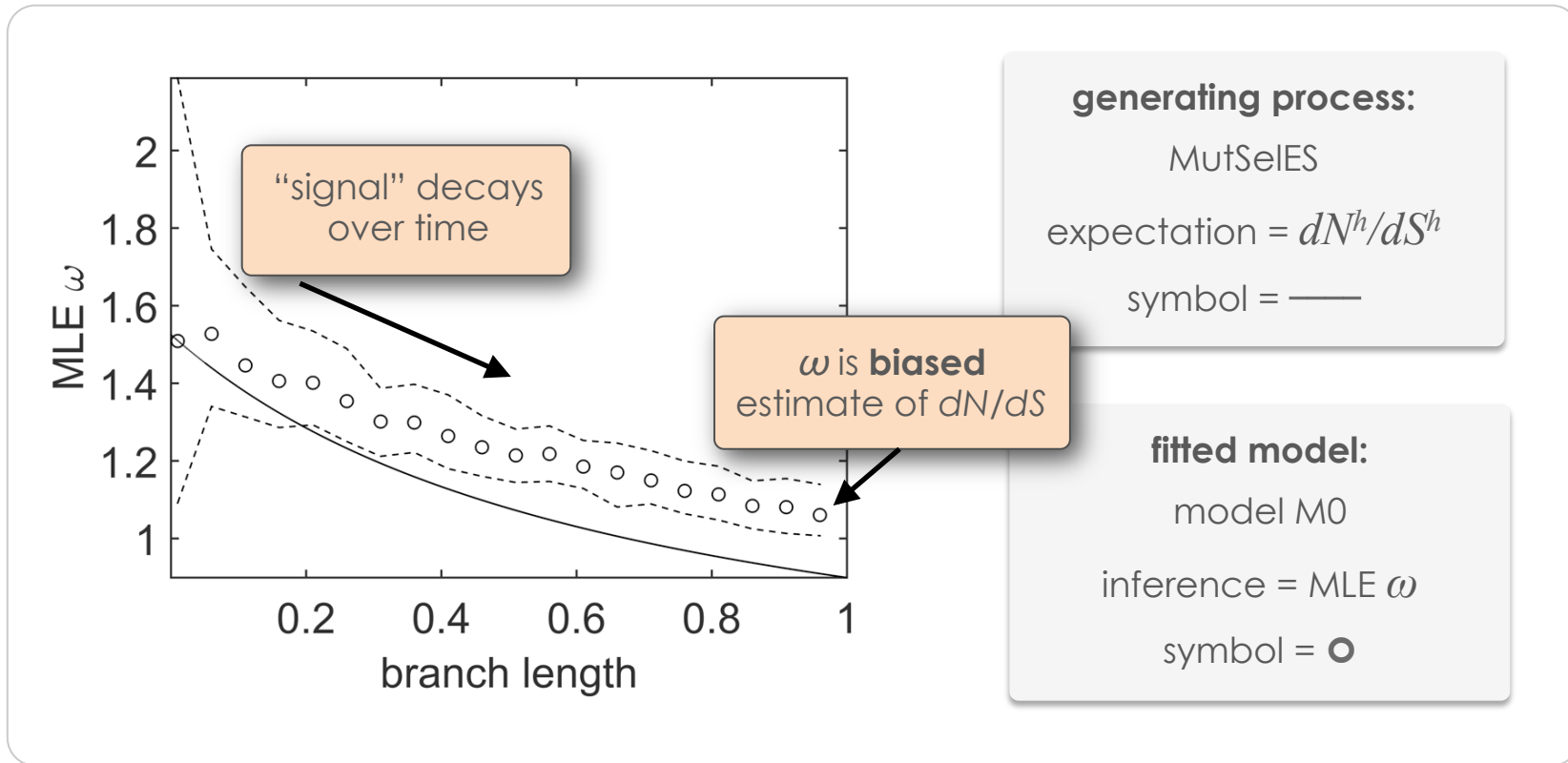
Mario dos Reis

Department of Genetics, Evolution and Environment, University College London, Gower Street, London WC1E 6BT, UK

First principles of population genetics are used to obtain formulae relating the non-synonymous to synonymous substitution rate ratio to the selection coefficients acting at codon sites in protein-coding genes. Two theoretical cases are discussed and two examples from real data (a chloroplast gene and a virus polymerase) are given. The formulae give much insight into the dynamics of non-synonymous substitutions and may inform the development of methods to detect adaptive evolution.

4. The non-synonymous rate during adaptive evolution

adaptive peak shift: MutSelES



conclusion : episodic models “work” because $w > 1$ is a consequence of a system moving towards a new fitness peak.

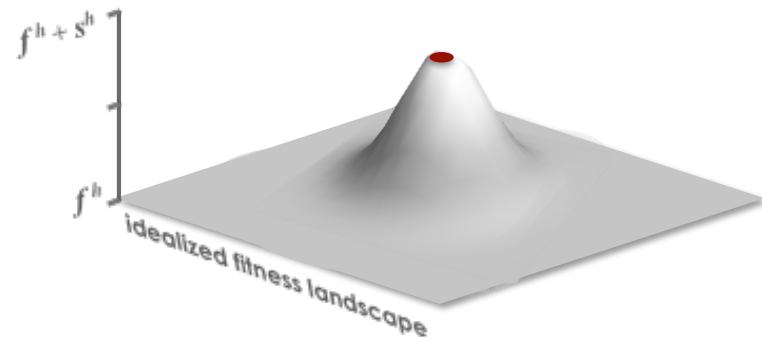
conclusion : episodic models “work” because they are sensitive to non-stationary behavior

Scenario 3: non-adaptive evolution

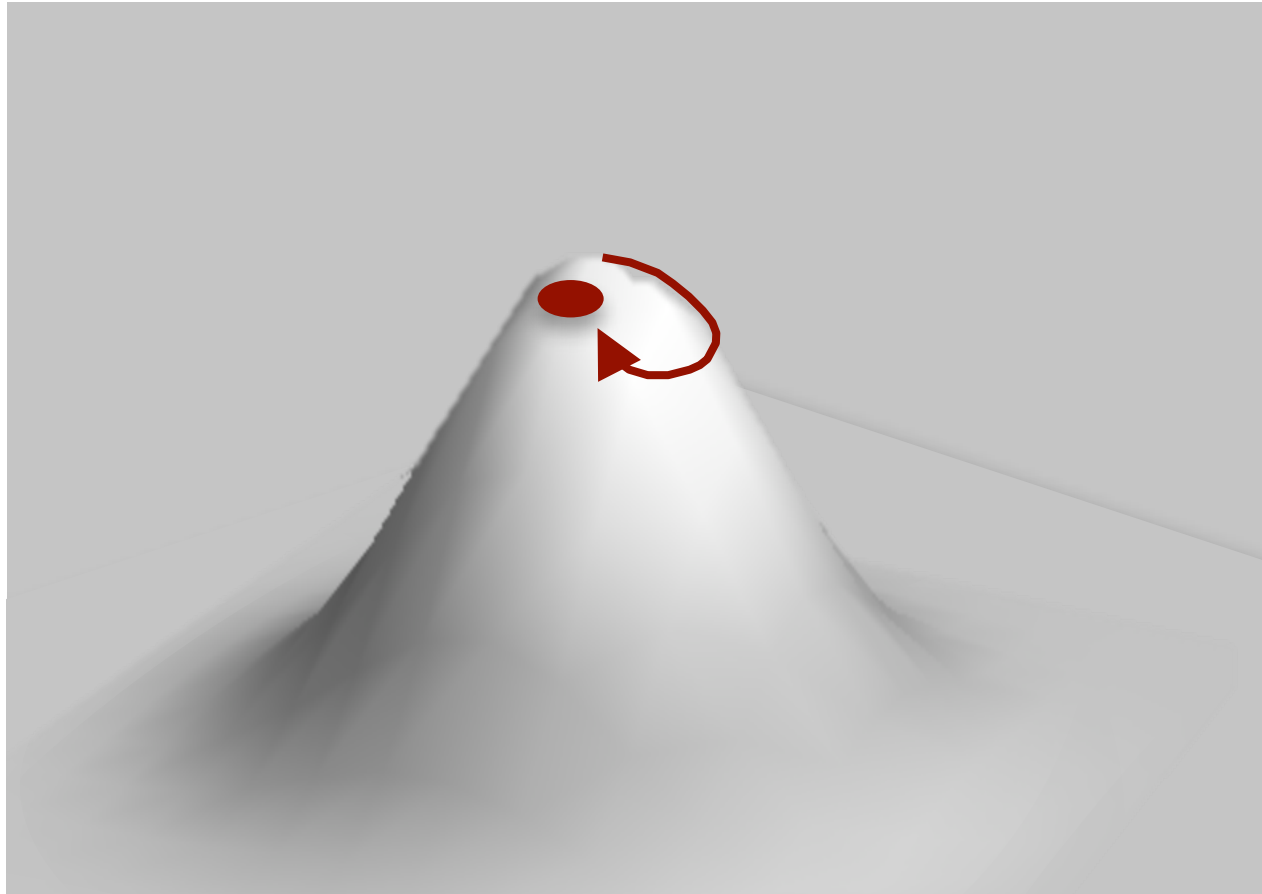
3. fitness coefficients are constant (fixed-peak)

Spielman and Wilke (2015)

- dN/dS must be ≤ 1 when fitness coefficients are fixed.
- positive selection is not possible on a stationary fitness peak

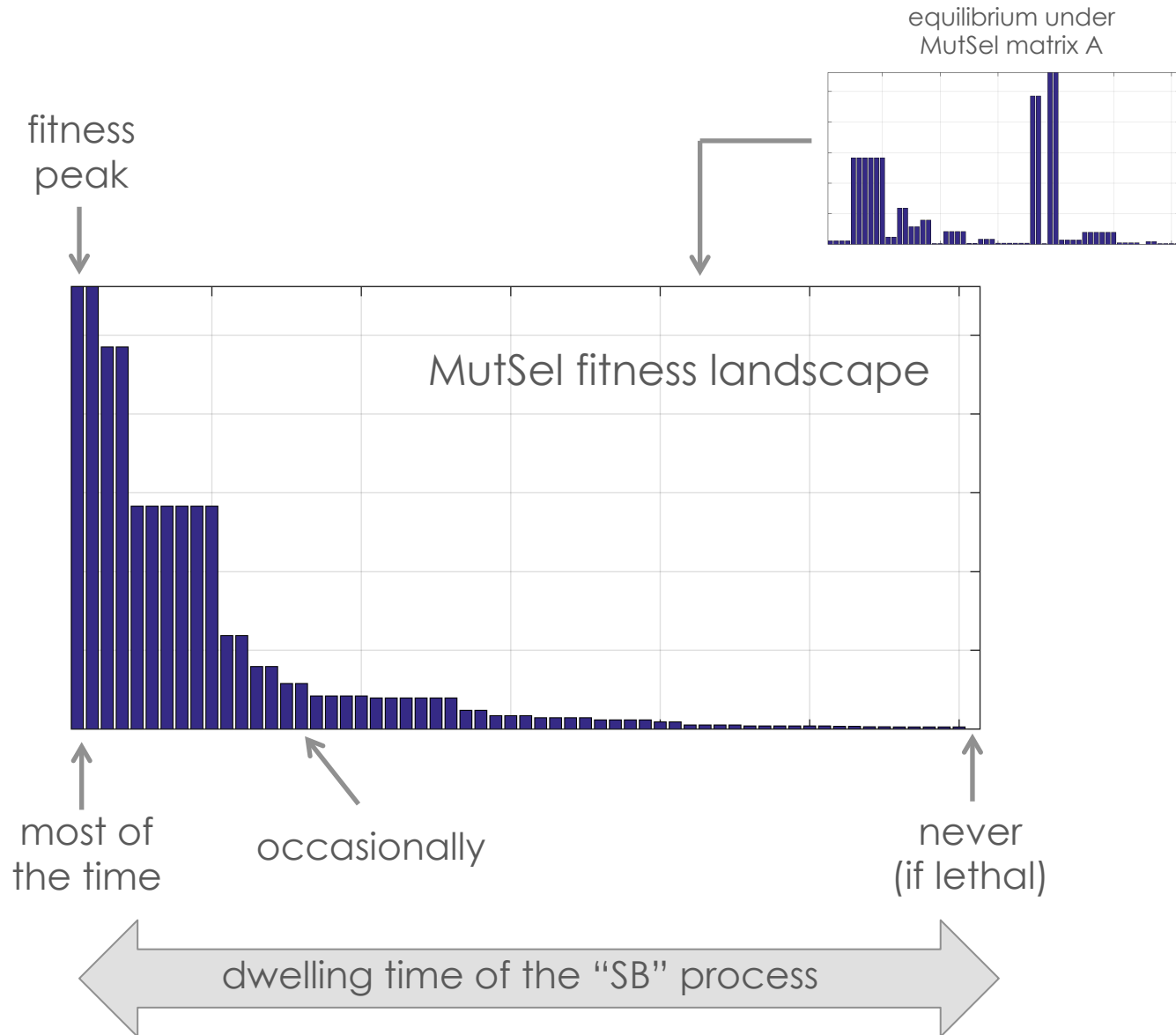


shifting balance: movement around peak

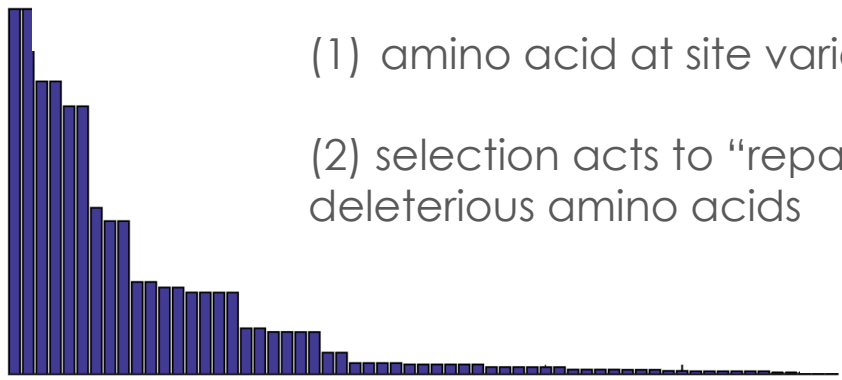


mutation and **drift** can move a pop. off a fitness peak

shifting balance: the MutSel landscape (Jones et al. 2016)



shifting balance: positive selection on a MutSel landscape



(1) amino acid at site varies over time

(2) selection acts to “repair” shifts to deleterious amino acids



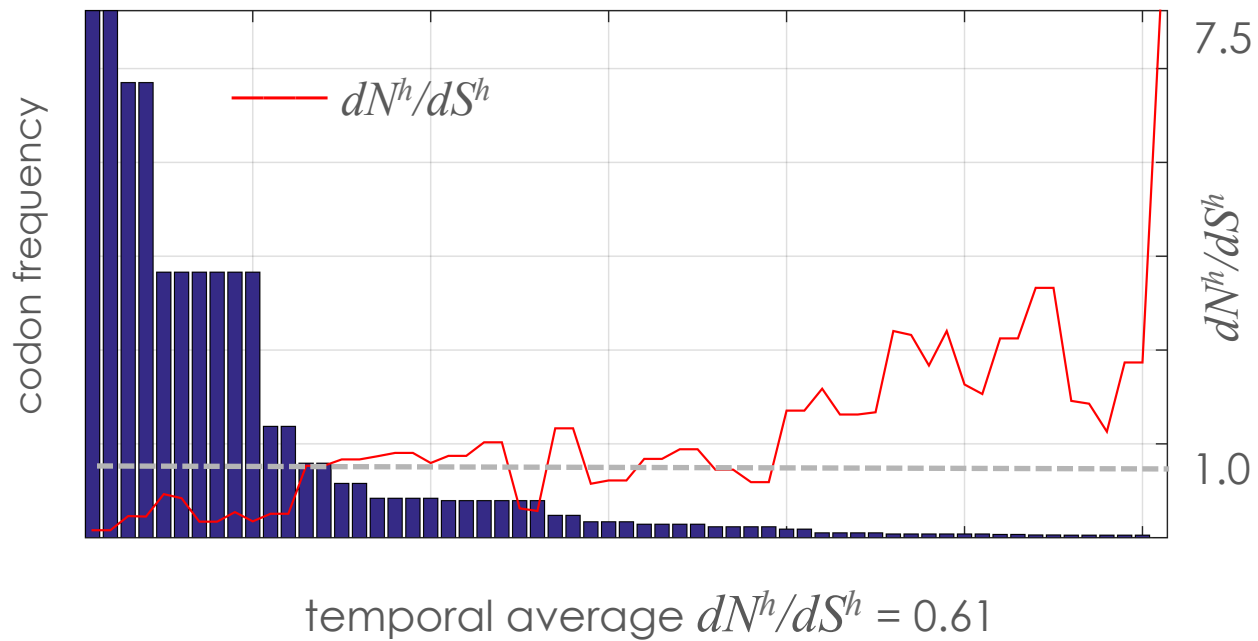
EXPECTED PROPORTION OF
MUTATIONS FIXED BY SELECTION

$$p_+^h = \frac{\sum_{(i,j)} \pi_i^h (A_{ij}^h - \mu_i) I_+}{\sum_{i \neq j} \pi_i^h A_{ij}^h}$$

conclusion: $p_+ > 0$ as long as number of viable amino acids > 1 at a site

shifting balance: the MutSel landscape

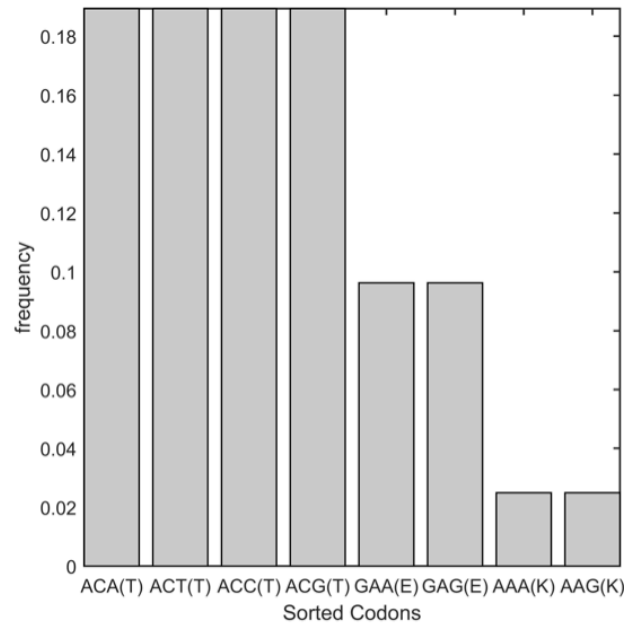
dN^h/dS^h depends on the current amino acid



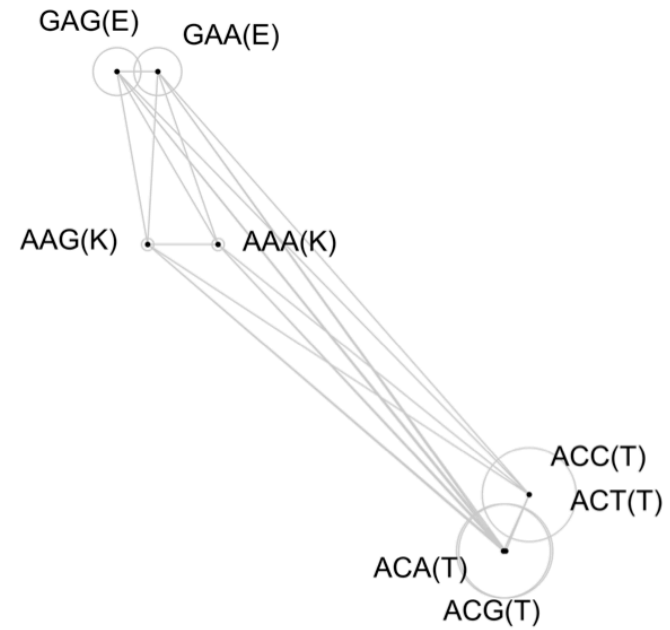
conclusion: positive selection operates on a stationary fitness peak in the same way as when there is an adaptive peak shift

landscapes have unique structures

MutSel landscape



McCandlish landscape

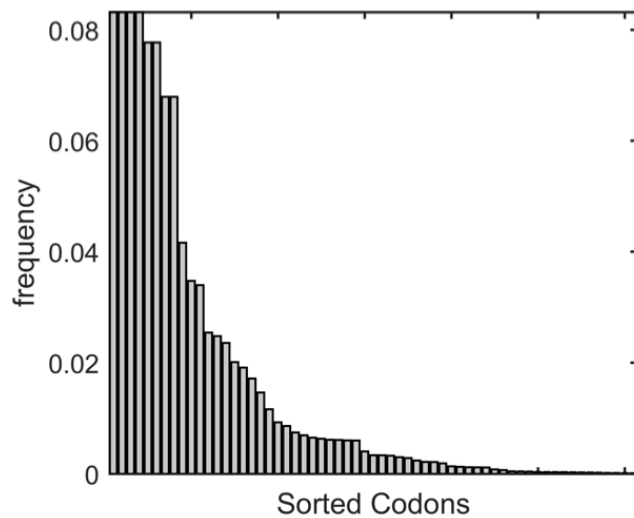


conclusion: A population can get to a sub-optimal codon (E) by drift and reside there for some time (b/c moving between T and E requires changes ≥ 2 codons).

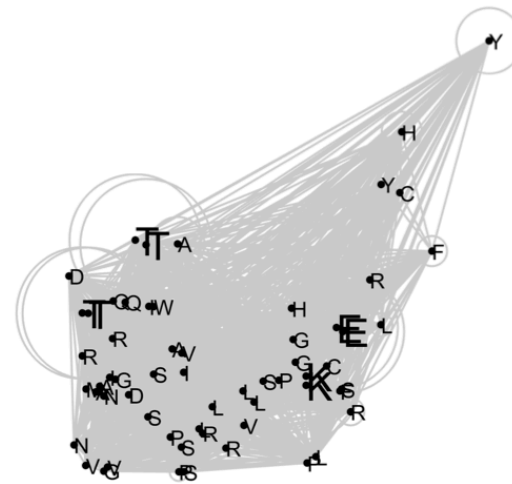
landscape structure depends on N

same site... 10x decrease in N (f^h have not changed!)

MutSel landscape



McCandlish landscape

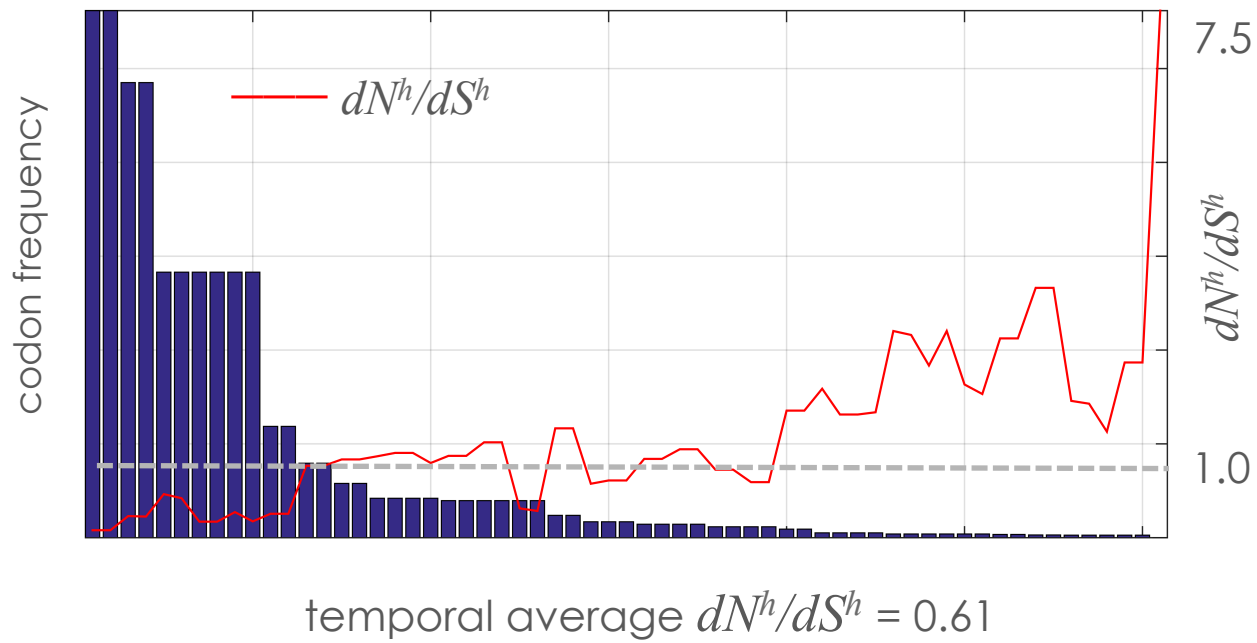


conclusion: decreasing N changes:

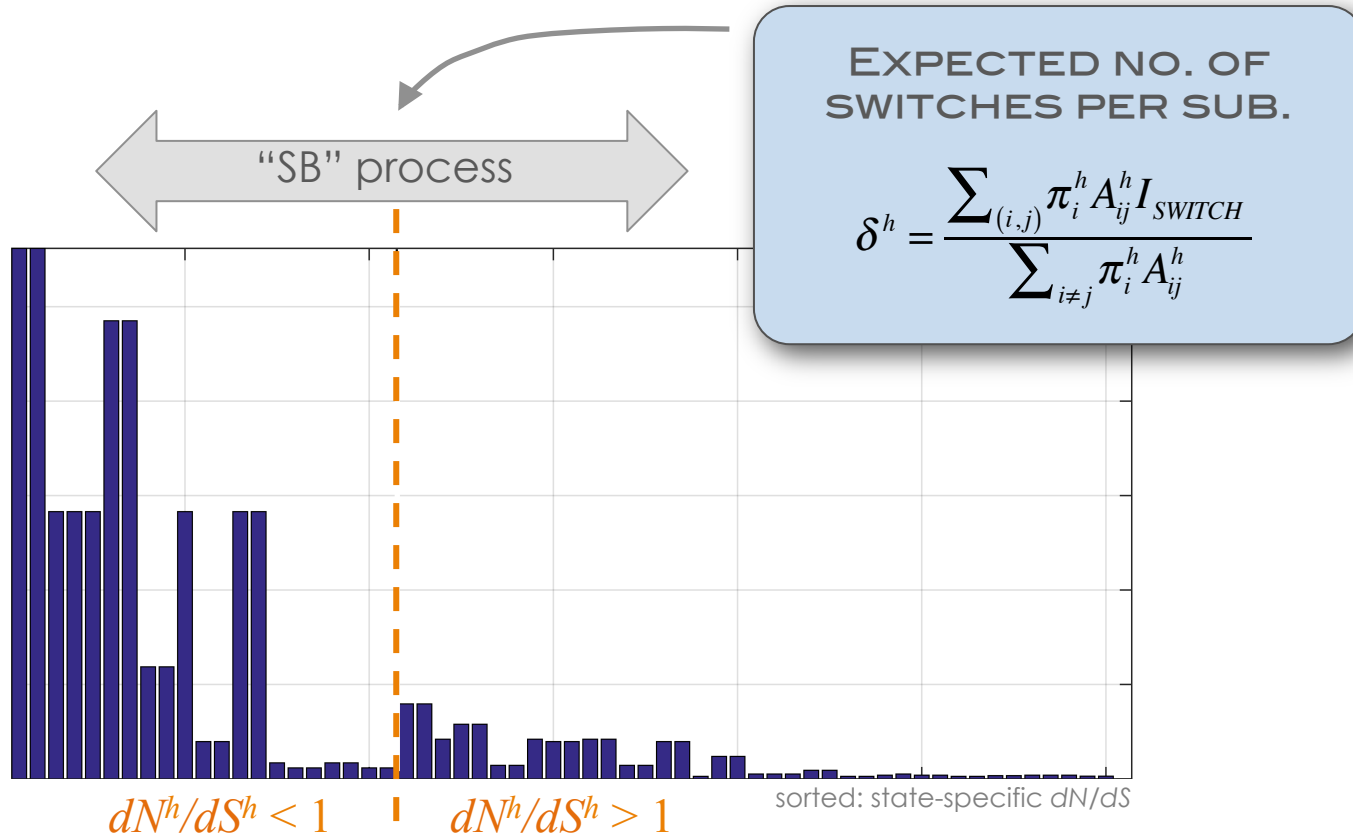
- i. the “space” for shifting balance
- ii. mean dN/dS
- iii. equilibrium frequencies

shifting balance: the MutSel landscape

dN^h/dS^h depends on the current amino acid



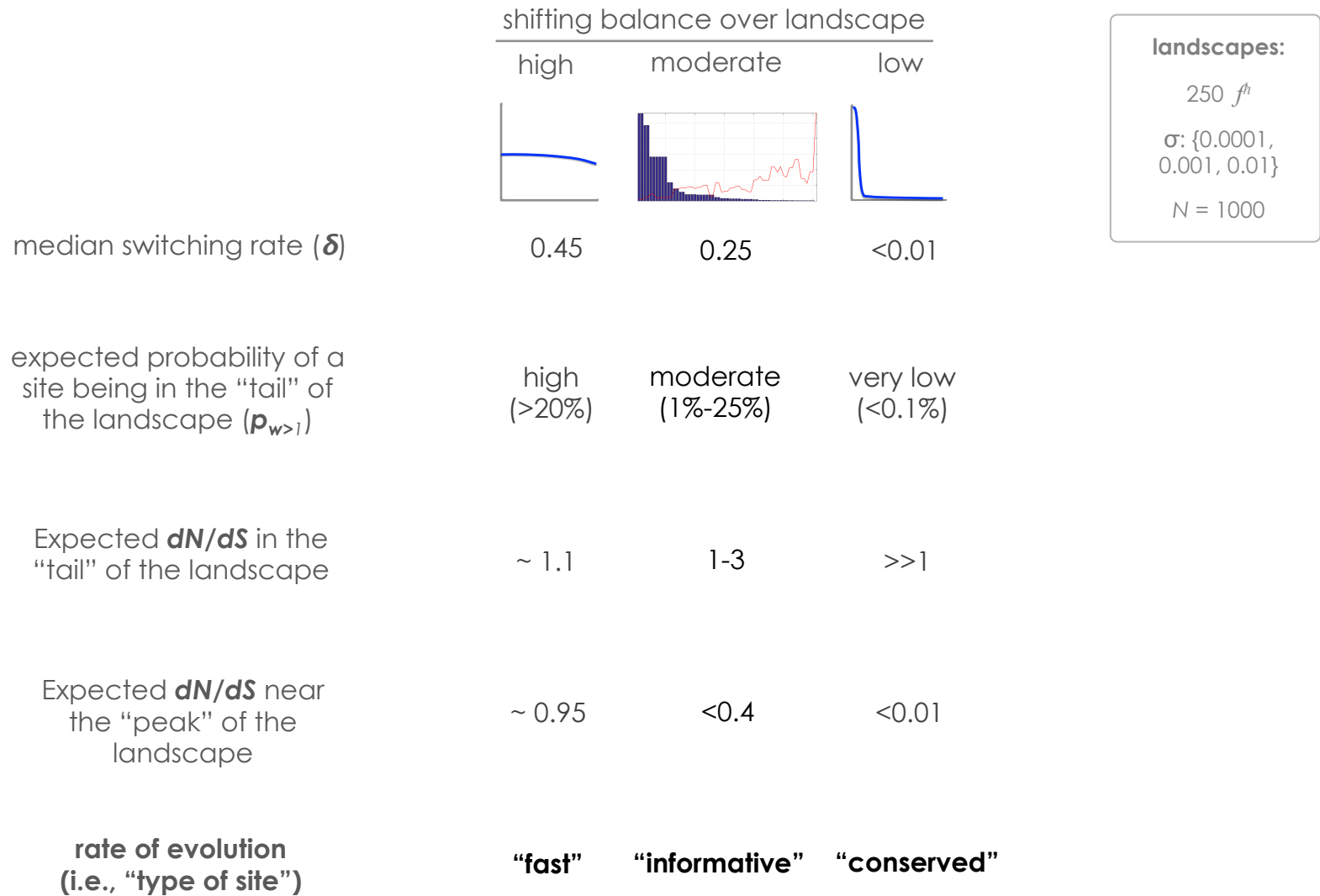
shifting balance: a mechanistic model



$$\omega^h < 1 = \frac{\sum_{i \in I_p^h} \frac{\pi_i^h}{p_1} A_{ij}^h I_N}{\sum_{i \in I_p^h} \frac{\pi_i^h}{p_1} \mu_{ij}^h I_N}$$

$$\omega^h > 1 = \frac{\sum_{i \in I_t^h} \frac{\pi_i^h}{p_2} A_{ij}^h I_N}{\sum_{i \in I_t^h} \frac{\pi_i^h}{p_2} \mu_{ij}^h I_N}$$

shifting balance: a mechanistic model



gene sequences

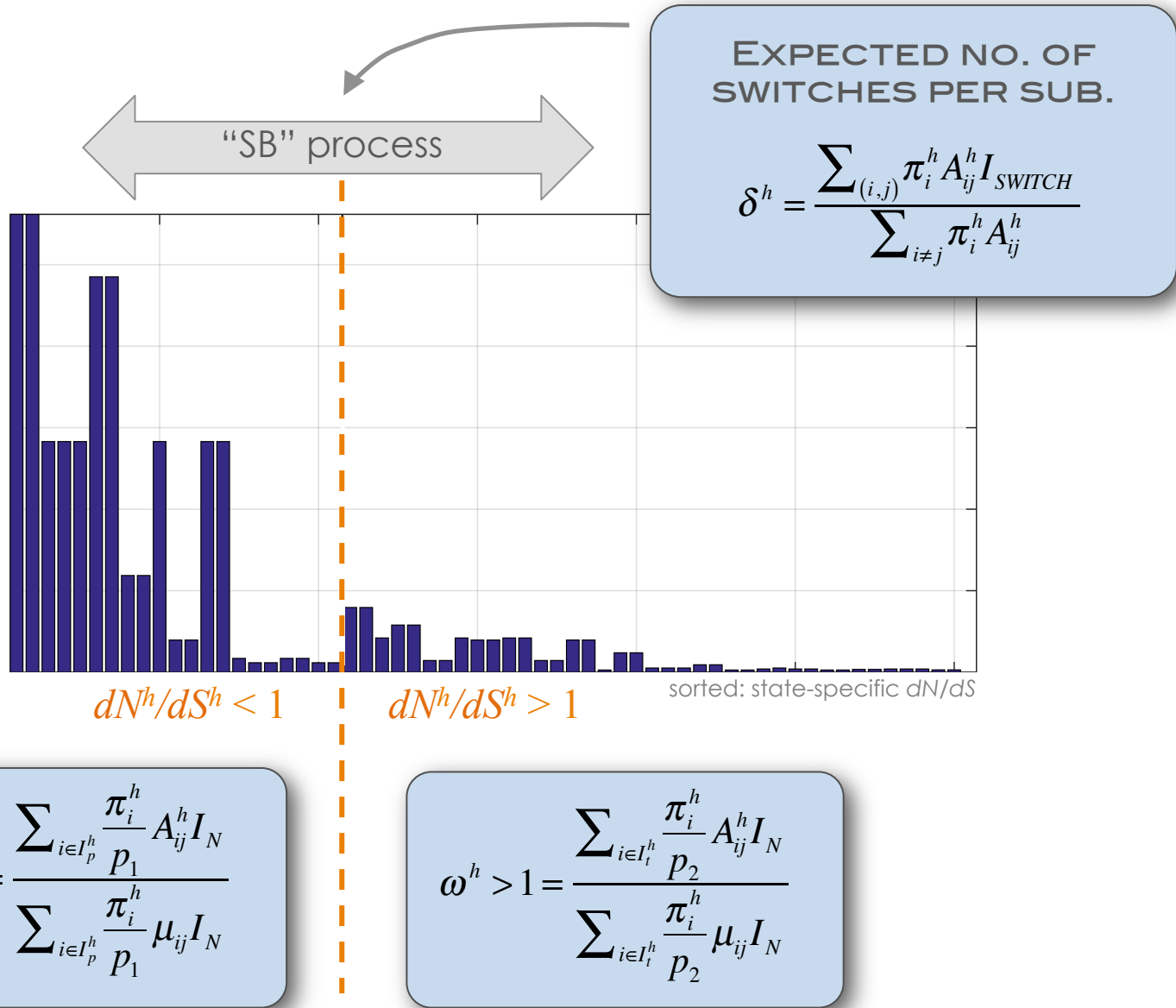
human
cow
rabbit
rat
opossum

GTG CTG TCT CCT GCC GAC AAG ACC AAC GTC AAG GCC GCC TGG GGC AAG GTT GGC GCG CAC
... .. G.C T.. ..TGC A..
...C ..T A.. ... A.TAA ... A.C ... AGC ...
... ..C ... G.A .ATA A.. ... AA. TG.G ... A.. ..T .GC ..T
... ..C ..G GA. ..TT C.. ..G ..A ... AT.TG ..A .GC ...

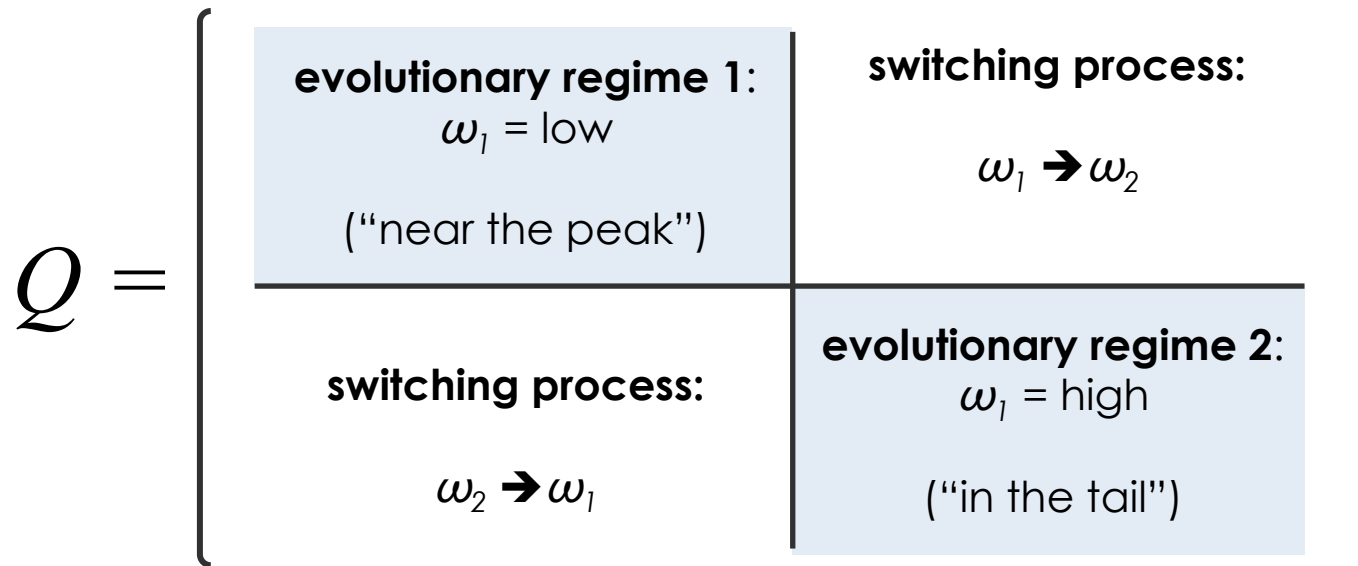
GCT GGC GAG TAT GGT GCG GAG GCC CTG GAG AGG ATG TTC CTG TCC TTC CCC ACC ACC AAG
... ..A .CTC ..AT AG.
.G.C ..C G.. T.. GG.
.G. ..T ..AC .A.A C.. GCT G..
..C ..T .CC ..C .CA ..T ..A ..T ..T .CC ..A .CCCTA

ACC TAC TTC CCG CAC TTC GAC CTG AGC CAC GGC TCT GCC CAG GTT AAG GGC CAC GGC AAG
...CGC G..
...C T.C .C.AG ... A.C ..A .C.
... T.T ... A.T ..T G.AC.CCT
..TC TC. .C.C A.C C.. ..T ..T ..T ...

shifting balance: a mechanistic model

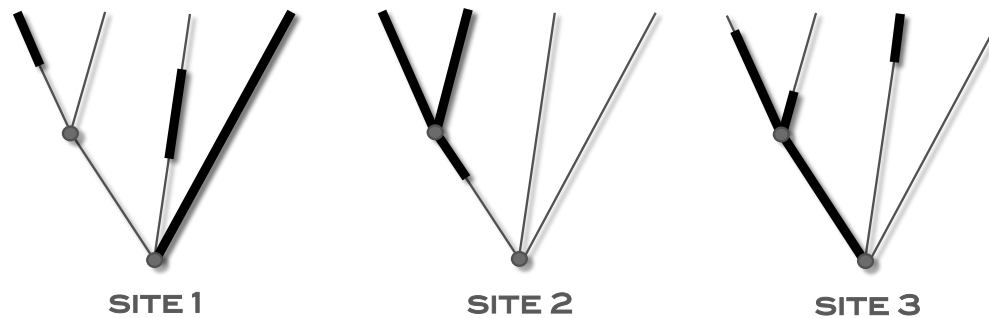


covariation-like model of evolution



covariation-like model of evolution
(phenomenological)

2 selective regimes (low & high): sites **CAN** switch regime

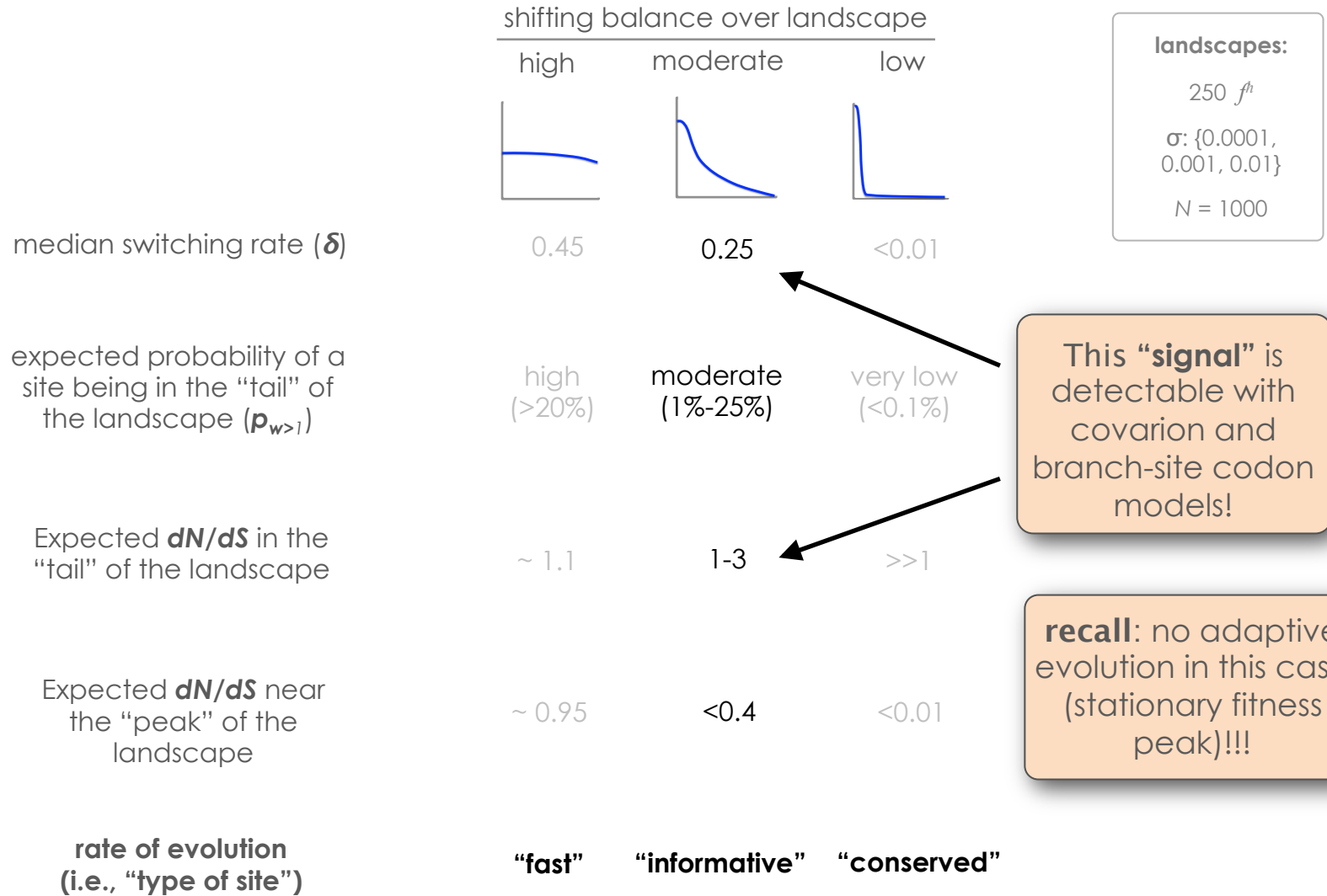


— **LOW** (ω_1) **—** **HIGH** (ω_2)

p_1 : proportion of time sites are in ω_1
 p_2 : proportion of time sites are in ω_2
switching: δ

the covariation-like codon model can be **fit to real data**

shifting balance: a mechanistic model



landscapes:
 250 f^i
 σ : {0.0001, 0.001, 0.01}
 N = 1000

summary

- standard codon models (single ω) assume frequency dependent selection, which yields a persistent $dN/dS > 1$
- episodic adaptive evolution leads to transient $dN/dS > 1$ (non-stationary process, with ω upwardly biased)
- MutSel landscapes can be complex and a site can reside at a sub-optimal state for extended periods of time
- protein evolution on a static fitness landscape has temporal dynamics that include positive selection
- rate variation among sites reflects the interplay between mutation, drift, and selection (i.e., shifting balance dynamics)

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